

# Growth and Structural Transformation\*

(Prepared for the Handbook of Economic Growth)

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## 1 Introduction

The one–sector growth model has become the workhorse of modern macroeconomics that is used for measuring aggregate economic activity and for addressing a wide range of important positive and normative issues. The popularity of the one–sector growth model is at least partly due to the fact that it captures in a minimalist fashion the essence of modern economic growth, which Kuznets (1973) in his Nobel prize lecture described as the sustained increase in productivity and living standards. By virtue of being a minimalist structure, the one–sector growth model necessarily abstracts from several features of the process of economic growth. One of the most important ones is structural transformation, that is, the reallocation of economic activity across agriculture, manufacturing and services.

Kuznets listed structural transformation as one of the six main features of modern economic growth. Structural transformation has also received a lot of attention in the policy debate of developed countries where various observers have claimed that the sectoral reallocation of economic activity is inefficient, and calls for government intervention. Understanding whether structural transformation arises as an efficient equilibrium outcome requires enriching the one–sector growth model to incorporate multiple sectors. More generally, this raises the question whether incorporating multiple sectors will sharpen or expand the insights that can be obtained from the one–sector growth model. Several researchers have recently begun to tackle these questions, and the objective of this chapter is to synthesize and evaluate their efforts.<sup>1</sup>

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<sup>1</sup>A different aspect of structural transformation that Kuznets also noted is the movement of the population from

A first step in the broad line of research on structural transformation is to develop extensions of the one–sector growth model that are consistent with the “stylized facts” of structural transformation. Accordingly, we begin this chapter by presenting the stylized facts of structural transformation and then we develop a multi–sector extension of the growth model that serves as a natural benchmark model to address the issue of structural transformation. Given the prominent role attributed to theories of balanced growth in the literature using the one–sector growth model, we start by asking whether it is possible to simultaneously deliver structural transformation and balanced growth. Recent work has identified several versions of the growth model that achieve this, and we present the results of this work in the context of our benchmark multi–sector model.

It turns out that the conditions under which one can simultaneously generate balanced growth and structural transformation are rather strict, and that under these conditions the multi–sector model is not able to account for the broad set of empirical regularities that characterize structural transformation. We therefore argue that the literature on structural transformation has possibly placed too much attention on requiring *exact* balanced growth, and that it would be better served by settling for *approximate* balanced growth instead. Put somewhat differently, we think that progress in building better models of structural transformation will come from focusing on the forces behind structural transformation without insisting on exact balanced growth. While the corresponding efforts to uncover the forces behind structural transformation are relatively recent, we describe some headway that has been made. We argue that the recent work suggests that the benchmark multi–sector model with approximate balanced growth is able to account for many salient features of structural transformation for the US, both qualitatively and quantitatively.

Armed with an extension of the one–sector growth model that incorporates structural transformation in an empirically reasonable fashion, we seek to answer the question of whether modeling structural transformation indeed delivers new or sharper insights into issues of interest. We argue that the answer to this question is yes, and we present several specific examples

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rural into urban areas, which goes along with the movement of employment out of agriculture. For an analysis of the rural–to–urban transformation in the context of the growth model, we refer the reader to the review paper of Greenwood and Seshadri (2005).

from the literature to illustrate this. These examples have in common that taking into account changes in the sectoral composition of the economy is crucial for understanding changes in aggregate outcomes. As we will see, this applies to important issues concerning economic development, regional income convergence, aggregate productivity trends, hours worked, business cycles, and wage inequality.<sup>2</sup>

## 2 The Stylized Facts of Structural Transformation

As mentioned in the introduction, structural transformation is defined as the reallocation of economic activity across three broad sectors (agriculture, manufacturing, and services) that accompanies the process of modern economic growth.<sup>3</sup> In this section, we present the stylized facts of structural transformation. While a sizeable literature on the topic already exists, including the notable early contributions of Clark (1957), Chenery (1960), Kuznets (1966), and Syrquin (1988).<sup>4</sup>, we think that improvements in the quality of previous data and the appearance of new data sets make it worthwhile for us to summarize the current state of evidence.

Because the process of structural transformation continues throughout development, it is desirable to document its properties using relatively long time series for individual countries. The early studies that we cited above attempted to do this. However, the authors of these studies typically had to piece together data from various sources, necessarily creating issues about the comparability of their results across time and countries. In addition, the time period for which data was available was still relatively short. Recent efforts by various researchers to reconstruct historical data have increased the availability of appropriate long time series data for the purposes of documenting structural transformation. Although one still has to piece together data from different sources to generate long time series for most countries, time coverage has

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<sup>2</sup>Matsuyama (2008) and Ray (2010) also offer recent review articles on structural transformation (or structural change, as Ray calls it). This chapter distinguishes itself from their reviews by the three-step approach that we take: document the stylized facts of structural transformation; build multi-sector extensions of the standard growth model that can account for them; employ these multi-sector extensions to improve our understanding of various aggregate phenomena.

<sup>3</sup>We follow much of the literature and use the term manufacturing in this context to refer to all activity that falls outside of agriculture and services. It might seem to be more appropriate to refer to this category as industry, because manufacturing is just the largest component of it, but we prefer to reserve the term “industry” to refer to a generic production category.

<sup>4</sup>The list of subsequent papers is too large for us to attempt to include it in its entirety.

improved and compatibility is much less of a problem than it was in the past. We are going to use the Historical National Accounts Database of the University of Groningen as our primary historical data source, which we complement with several other data sources to increase the length of the periods covered.<sup>5</sup>

While it is conceptually desirable to examine changes for individual countries over long time series, and there is now more opportunity to do so, limiting attention to individual countries narrows the perspective unnecessarily. To begin with, it effectively restricts the set of countries that can be studied to those that are currently rich, and so it leaves open the question of whether currently poor countries show the same regularities that currently rich countries showed when they were poor a century or two ago. Limiting attention to long time series data has the additional disadvantage that despite major improvements in constructing historical time series, they typically do not reach the quality of the best data sets for recent years. Therefore, we document the regularities of structural transformation also for five data sets that cover a relatively large set of developing and developed countries during the last thirty or so years: the Benchmark Studies of the International Comparisons Program as reported by the Penn World Table (PWT), EUKLEMS, the National Accounts from the United Nations Statistics Division, the OECD Consumption Expenditure Data, and the World Development Indicators (WDI).<sup>6</sup>

## **2.1 Measures of Structural Transformation**

Before presenting any data, it is useful to briefly note some aspects of measuring economic development and structural transformation.

The two most common measures of economic development at the aggregate level are GDP per capita and some measure of productivity (typically GDP per worker or GDP per hour, depending upon data availability), each expressed in international dollars. While these two measures often coincide, there are cases in which they differ. For example, several European

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<sup>5</sup>Appendix A contains a detailed description about the historical data sources that we use. Many of them are also underlying the recent historical studies by Dennis and Iscan (2009) about structural transformation in the United States and by Alvarez–Cuadrado and Poschke (2011) about structural transformation in twelve industrialized countries including the United States.

<sup>6</sup>We again refer the reader to Appendix A for the details regarding the data sets and how we use them to construct measures of economic activity at the sector level.

economies have similar values of GDP per hour as the US, but GDP per capita can be as much as 25 percent lower than in the US because hours per adult are much lower. Without knowing the exact context of the issue being addressed, it is unclear whether one should categorize these European countries as equally or less developed than the United States.

Having raised this issue, in this chapter we choose to always measure the level of development by GDP per capita in 1990 international \$. Three reasons motivate this choice. First, in order to be able to identify threshold effects and the like, we insist on the comparability of the GDP numbers across different data sets, and GDP per capita is the only measure that is available for most countries and most of the time. Second, the standard models of structural transformation take labor supply to be exogenous, implying that they abstract from differences in hours worked. Third, since some of the models that we will consider emphasize the role of income effects for structural transformation, it seems to be appropriate to characterize the patterns of sectoral reallocation conditional on income. Irrespective of these three reasons for using GDP per capita, we emphasize that most of our figures would look similar if instead we used one of the productivity measures when they are available.

Next we turn to measuring structural transformation. The three most common measures of economic activity at the sectoral level are employment shares, value added shares, and final consumption expenditure shares. Employment shares are calculated either by using workers or hours worked by sector, depending on data availability. Value added shares and final consumption expenditure shares are typically expressed in current prices (“nominal shares”), but they may be expressed in constant prices also (“real shares”). While there is a tendency in the literature to view the different measures as interchangeable when documenting how economic activity is reallocated across sectors over time, one of the issues that we want to emphasize in this chapter is that they are in fact distinct. In particular, as we will discuss in detail later on, it is critical to be aware of the distinctions among the different measures when doing quantitative work because even when they display the same qualitative behavior, the quantitative implications may be quite different. Moreover, there are some striking cases in which they display differences even in the qualitative behavior.

Probably the most important reason for the differences between the measures of structural

transformation is that employment shares and value added shares are related to production whereas final consumption expenditure shares are related to consumption. There are two main reasons why production and consumption measures may display different behavior. The first reason is that final goods are not only consumed but also invested. While this point is particularly relevant in open economies with investment opportunities in other countries, it applies also to closed economies. The second reason why production and consumption measures may display different behavior is that consumption expenditure refer only to expenditure on final goods whereas production value added does not distinguish whether the value added represents a final good or an intermediate good.

A simple example will help to illustrate the distinction between final goods and intermediate goods that is relevant here. Consider the purchase of a cotton shirt from a retail establishment. Because the cotton shirt is a “good” as opposed to a “service”, in terms of final consumption expenditure, the entire expenditure will be measured as final consumption expenditure of the manufacturing sector. However, in terms of value added in production, the same purchase will be broken down into three pieces: a component from the agricultural sector (i.e., the cotton that was used in making the shirt), a component from the manufacturing sector (i.e., the processing of the cotton and the production of the shirt), and a component from the service sector (i.e., the distribution and retail trade services where the shirt was purchased).

The end result of this is that although the same sectoral labels are used when disaggregating GDP into final expenditure and value added, the resulting measures of sectoral shares are conceptually distinct. It follows that both quantities and prices may differ between final consumption and production value added, implying that there is no reason to expect the implied shares to exhibit similar behavior. This will be of particular relevance when connecting models of structural transformation to the data, which we will discuss in detail below.

The previous discussion emphasized the difference between production and consumption measures. However, even two measures that focus on production might contain different information. One example comes from Kuznets (1966), who showed for the early part of US development that the employment share of services increased considerably at the same time that the value added share of services remained almost constant.



Having emphasized that each of the three measures of economic activity at the sectoral level is distinct, we also want to note that each of them has its limitations as a singular measure. For the case of sectoral employment shares, a key issue is that employment may not reflect changes in “true” labor input, for example, because there are systematic differences in hours worked or in human capital per worker across sectors that vary with the level of development. For the case of value added and consumption expenditure shares, a key issue arises from the need to distinguish between changes in quantities and prices. This is often difficult empirically because reliable data on relative price comparisons across countries are hard to come by. And, as noted previously, in an open economy, we know that consumption and production need not coincide, so that neither measure alone is sufficient.

## **2.2 Production Measures of Structural Transformation**

In this subsection we document the patterns of structural transformation based on examining production measures in several different data sets. We first review the available historical time series evidence for currently rich economies. We then turn to the evidence for currently rich and poor countries.

### **2.2.1 Evidence from Long Time Series for Currently Rich Countries**

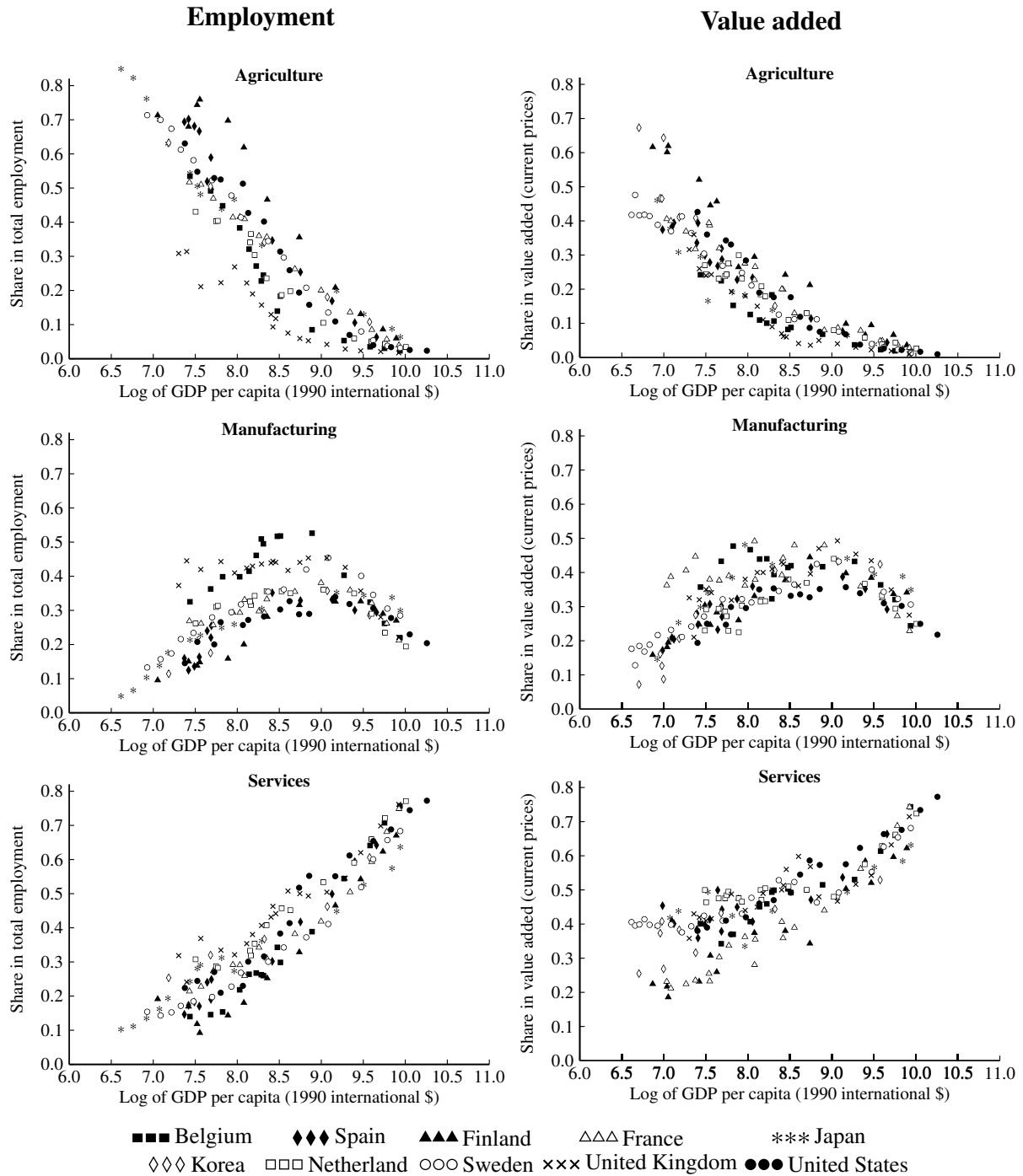
We construct individual time series of sectoral employment shares and value added shares over the 19th and 20th century for the following ten countries: Belgium, Finland, France, Japan, Korea, Netherlands, Spain, Sweden, United Kingdom and United States.<sup>7</sup> Since the early data is sketchy and we want to highlight trends over long periods of time, we report the latest available observation for each decade, if any. We note that for these historical time series we only have measures based on the production.

Figure 1 plots the historical time series. The vertical axis is either the share of employment or the share of value added in current prices in the three broad sectors of interest. The horizontal axis is the log of GDP per capita in international dollars as reported by Maddison. The figures clearly reveal what the literature views as the stylized facts of structural transformation.

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<sup>7</sup>For a detailed description of the data sources, see the Appendix A.

**Figure 1: Sectoral Shares of Employment and Value Added – Selected Developed Countries 1800–2000**



**Source:** Various historical statistics, see Appendix A.

Specifically, over the last two centuries, increases in GDP per capita have been associated with decreases in both the employment share and the nominal value added share in agriculture, and increases in both the employment share and the nominal value added share in services. Manufacturing has behaved differently from the other two sectors: its employment and nominal value added shares follow an inverted U shape, that is, they are increasing for lower levels of development and decreasing for higher levels of development.

Figure 1 reveals two additional regularities that have been somewhat less appreciated in the context of structural transformation. First, focusing on the agricultural sector, we can see that for low levels of development the value added share is considerably lower than the employment share. This finding is interesting in light of the fact that countries which are currently poor tend to have most of their workers in agriculture although agriculture is the least productive sector.<sup>8</sup> Second, focusing on the service sector, we see that both the employment share and the nominal value added share for the service sector are bounded away from zero even at very low levels of development; the lowest value added share of services is around 20% and the lowest employment share is around 10%.<sup>9</sup> Third, the figure for the nominal value added share in services suggests that for most countries there is an acceleration in the rate of increase when the log of GDP per worker reaches around 9.<sup>10</sup> Inspecting the graphs for the other two nominal value added shares more closely, we also note that the nominal value added share for manufacturing peaks around the same log GDP at which the nominal value added share for the service sector accelerates, and so it appears that the accelerated increase in the value added share of services coincides with the onset of the decrease in the nominal value added share for manufacturing sector.

### **2.2.2 Evidence from Recent Panels for Currently Rich and Poor Countries**

We now turn to an examination of production measures from several more recent data sets, which compared to the historical data are of higher quality and include also countries that are currently poor as well as additional variables (nominal versus real, hours versus employment).

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<sup>8</sup>See Caselli (2005) and Restuccia et al. (2006) for evidence on this point.

<sup>9</sup>This finding is confirmed by the historical study of Broadberry et al. (2011), who present evidence for England during the 14<sup>th</sup> century that the employment share of services was around 20%.

<sup>10</sup>See Buera and Kaboski (2011) for additional evidence on this point in a much larger cross section of countries.

The goal of the subsection is to assess the stylized facts of structural transformation that we documented for the historical data, as well as to take advantage of the richer data available so as to examine additional dimensions of structural transformation.

### **Evidence from EUKLEMS**

We start with EUKLEMS, which is compiled at the Groningen Growth and Development Center. EUKLEMS' primary strength in documenting patterns in employment and value added shares is that it has the most complete information for all variables of interest, including sectoral hours worked, and that its value added data have been constructed from the national accounts of individual countries following a harmonized procedure that aims to ensure cross-country comparability (for example, a common industry classification was used and price indices were constructed in a similar way across countries).<sup>11</sup> EUKLEMS' primary weakness is that its coverage is limited to countries with relatively high income; Korea during the early 1970s is the country with the lowest income in the sample.

We first document the evolution of the shares of sectoral hours worked and nominal value added as functions of the level of development for five non-European countries – i.e., Australia, Canada, Japan, Korea and the United States, as well as for the aggregate of fifteen EU countries.<sup>12</sup> The data are plotted in Figures 2. The vertical axis is either the share of total hours worked or the share of value added in current prices in the three broad sectors of interest. As before, the horizontal axis is the log of GDP per capita in international dollars from Maddison.

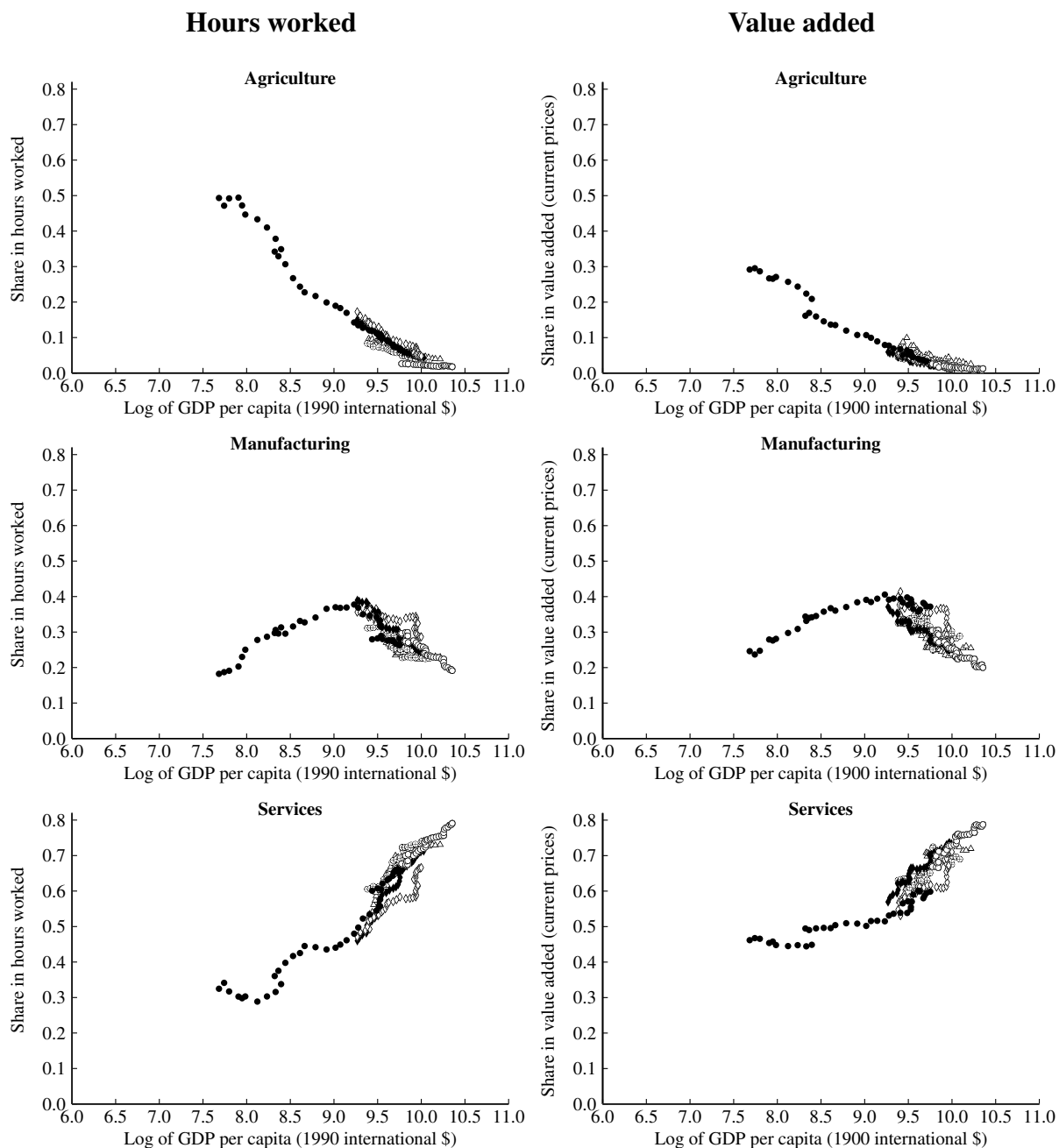
The plots in Figure 2 confirm several patterns from the historical times series. First, the shares of hours worked and nominal value added for agriculture tend to decrease with the level of development for all countries, whereas the shares for services tend to increase with the level of development for all countries. Second, taken as a whole, the data are consistent with an inverted U shape for the shares in the manufacturing sector, although all countries except for Korea have decreasing manufacturing shares. Third, the series for both shares as a function of GDP per capita are quite consistent across countries. That is, not only are the qualitative

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<sup>11</sup>For more detail see O'Mahony and Timmer (2009) and Timmer et al. (2010).

<sup>12</sup>These are Austria, Belgium, Denmark, Finland, France, Germany, Greece, Ireland, Italy, Luxembourg, Netherland, Portugal, Spain, Sweden, and the United Kingdom.

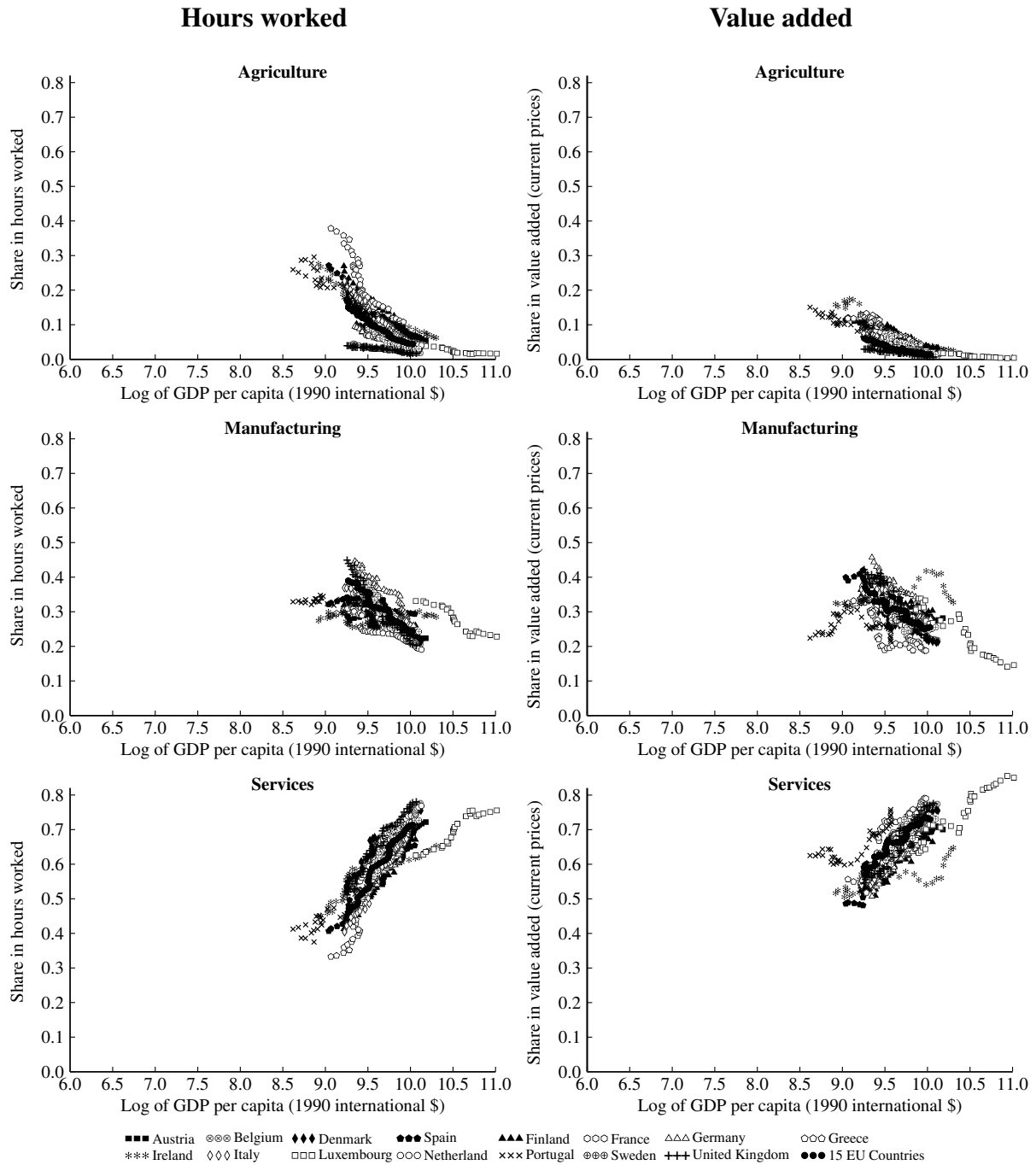
**Figure 2: Sectoral Shares of Hours Worked and Nominal Value Added – 5 Non-EU Countries and Aggregate of 15 EU Countries from EUKLEMS 1970–2007**



△△△ Australia ⊕⊕⊕ Canada ◆◆◆ 15 EU Countries ◇◇◇ Japan ●●● Korea ○○○ United States

**Source:** EUKLEMS, PWT6.3

**Figure 3: Sectoral Shares of Hours Worked and Nominal Value Added – 15 EU countries from EUKLEMS 1970–2007**



Source: EUKLEMS, PWT6.3

patterns very similar, but so too are the quantitative patterns. This is of particular interest given the considerable attention that has been placed on the role of openness in the growth miracle of Korea (Korea liberalized its manufacturing trade starting in the 1960s and became one of the most open countries in the world). Although to a lesser extent, one could make similar statements for the case of Japan.

Although this last finding might tempt one to conclude that openness is not a quantitatively important determinant of sectoral allocations and structural transformation, we do want to caution the reader against jumping too quickly to this conclusion. Figure 3 shows the same series separately for the 15 EU countries. Although all countries display the same qualitative patterns, there is now substantial heterogeneity in the cross section at any given level of development. This is at least consistent with the view that these countries form a fairly integrated free-trade zone, thereby allowing for a high degree of specialization, and significant differences in how economic activity is allocated across broad sectors.<sup>13</sup>

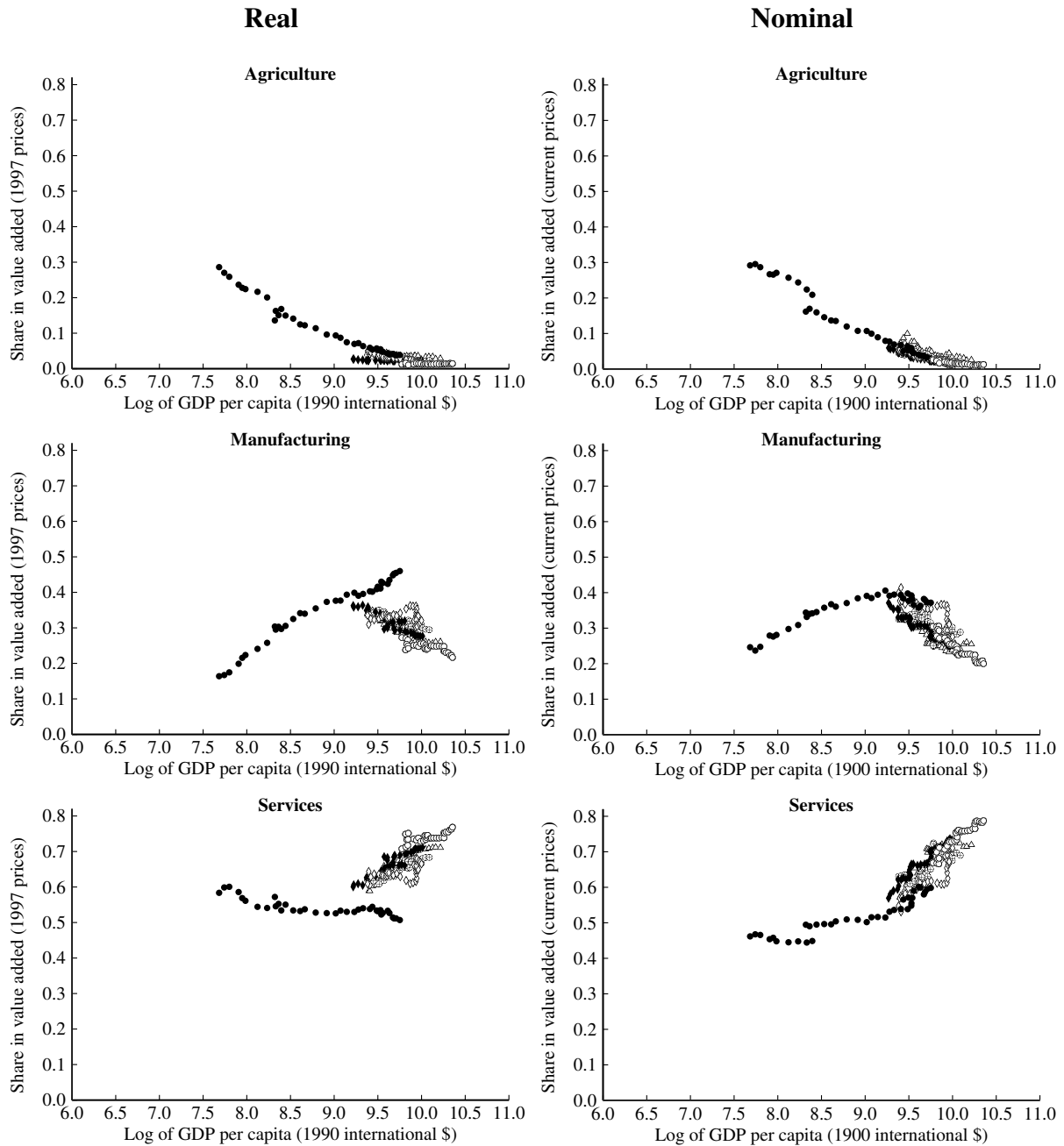
Next, we next turn our attention to possible differences between real and nominal shares of sectoral value added, where nominal refers to current prices and real refers to constant prices. Kuznets (1966) concluded that the early available data showed similar qualitative patterns for nominal and real shares. We revisit this comparison because EUKLEMS has more recent and higher quality data than were available to Kuznets. Figure 4 plots the real shares of sectoral value added in the left panel and, for comparison, the nominal shares from Figure 2 in the right panel. The plots show that the qualitative patterns of real and nominal value added shares are fairly similar to each other, confirming what Kuznets found for the earlier data.

One important exception is Korea where the manufacturing share rose to half of real value added, which is considerably higher than in the other countries on the graph. At the same time, the manufacturing share of nominal value added flattened out around the maximum share for the other countries. Moreover, the real service share remained below the service share of the other countries, and actually fell somewhat. At the same time, the nominal service share stayed mostly flat. These observations imply that the price of manufacturing relative to total value added fell by more in Korea than in the other countries. This is consistent with the view

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<sup>13</sup>Some of the series that we consider later on in this section will reveal differences between Korea and the other countries.

**Figure 4: Sectoral Shares of Real and Nominal Value Added –  
5 Non-EU Countries and Aggregate of 15 EU Countries from EUKLEMS 1970–2007**



△△△ Australia ⊕⊕⊕ Canada ◆◆◆ 15 EU Countries ◇◇◇ Japan ●●● Korea ○○○ United States

**Source:** EUKLEMS, PWT6.3



that during Korea's massive trade liberalization the relative price of manufactured goods fell considerably at the same time as the real growth rate of manufacturing increased considerably.

### **Evidence from the WDI and the UN Statistics Division**

As previously noted, the main shortcoming of both the historical data and of EUKLEMS is that the coverage is limited to countries that have fairly high income today. It is therefore of interest to verify whether the stylized facts of structural transformation extend to data sets that cover countries that are poor today. The two obvious data sets to use for this are the World Development Indicators (WDI) and the National Accounts that the United Nations Statistics Division collects.

We use the WDI for employment by sector, which it reports since 1980 based on the data published by the International Labor Organization (ILO). We emphasize that these data are about employed workers instead of hours worked and are of considerably lower quality than EUKLEMS because there is much less harmonization underlying the construction of the data. This leads to comparability issues of WDI data across countries. Moreover, these data are not uniformly available over time for all countries. We use the national accounts of the United Nations Statistics Division for value added by sector. Unlike the WDI, the UN Statistics Division provides continuous coverage for a large number of countries between 1970 and 2007 and makes an explicit effort to harmonize the national accounts data so as to ensure that they are comparable across different countries.

Figure 5 plots sectoral employment shares from the WDI against GDP per capita from Maddison. The plots confirm that in terms of sectoral employment shares the basic qualitative regularities of structural transformation also hold outside the set of rich countries for which EUKLEMS has data. Specifically, it is the case again that the agricultural employment share decreases in the level of development and that the employment share of services increases in the level of development. Moreover, the employment share in manufacturing is strongly increasing at lower levels of development (log of GDP per worker smaller than 9.5) before flattening out and then decreasing somewhat for higher levels of development. While this pattern is consistent with an inverted U shape, we note that the downward sloping part is not very pronounced in the

WDI data.

Not surprisingly, the plots also show that employment shares do take on much more extreme values than can be found in EUKLEMS. For example, now the employment share of agriculture can be as high as 70% percent and the employment shares of manufacturing and services can be as low as only 10%. Lastly, for a given level of development the plots show much greater variability in the employment shares relative to what we found in the EUKLEMS data. The extent to which this simply reflects greater measurement error due to lack of comparability and other factors is an open question.<sup>14</sup>

Figure 6 plots nominal value added shares by sector from the UN Statistics Division against GDP per capita from Maddison. Since these data have complete coverage for many rich and poor countries, they come close to a balanced panel. We therefore also plot the fitted nominal value added shares from panel regressions. We emphasize that the regressions are presented simply as a way of summarizing some patterns in the data and not as a way of testing any theory. For each sector we regress nominal value added shares on country fixed effects and the level, square, and cube of GDP per worker.<sup>15</sup> We include countries for which no observations are missing, that were not communist, and that had more than a million of inhabitants during 1970–2007. Details regarding the construction of the panel of countries and the regression results are contained in Appendix B.

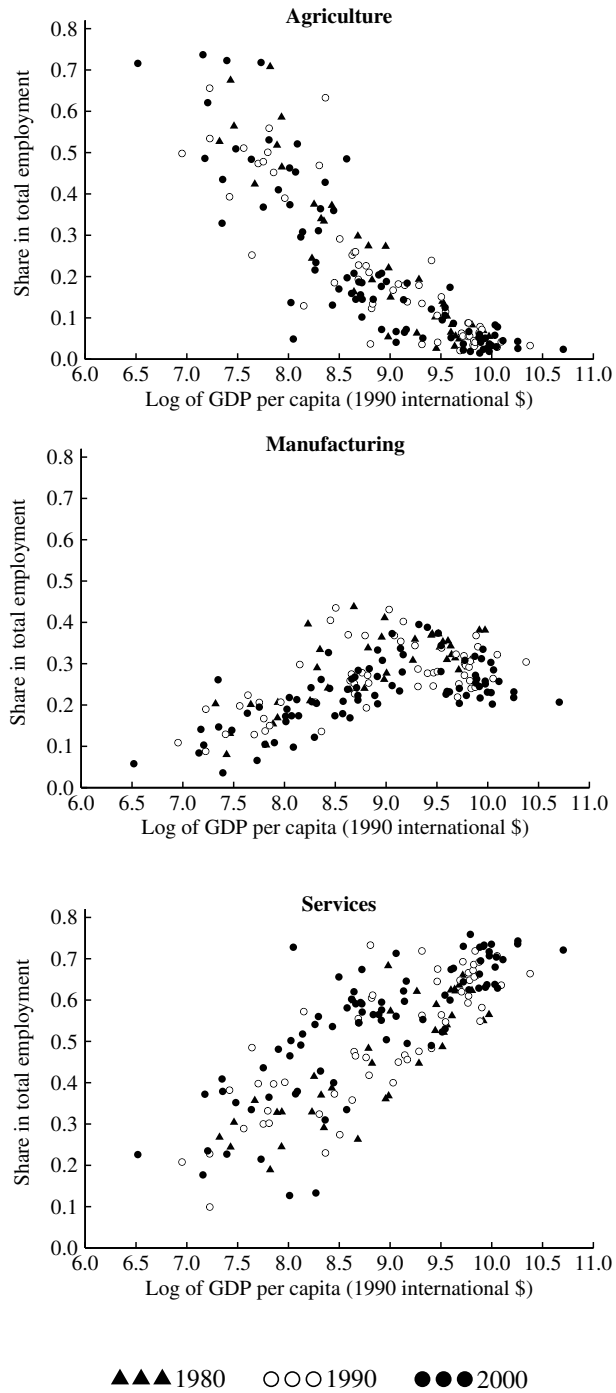
The fitted curves reveal the same qualitative patterns that we have documented previously. It is of particular interest that the inverted U shape clearly emerges for manufacturing value added. Moreover, it is of interest that the fitted curve for services indicates an acceleration of the service share when the log of GDP per capita reaches a threshold value around 9 and the share of manufacturing value added peaks. Interestingly, this feature occurs at a similar threshold value also for the historical time series which we discussed above.

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<sup>14</sup>See Herrendorf and Schoellman (2011) for further discussion of the role that measurement error can play even in the United States.

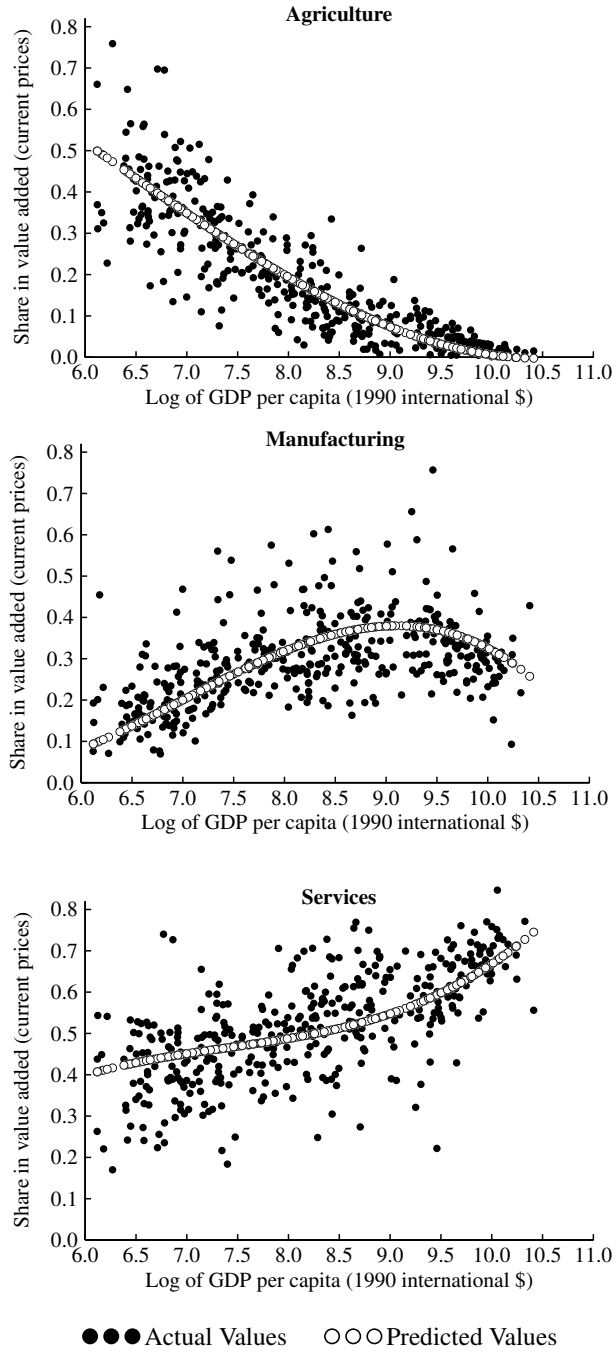
<sup>15</sup>We report results for a cubic polynomial since adding higher order terms did not have a significant effect on the fitted relationships.

**Figure 5: Sectoral Shares of Employment – Cross Sections from the WDI 1980–2000**



**Source:** World Development Indicators 2010

**Figure 6: Sectoral Shares of Nominal Value Added –  
Cross Sections from UN National Accounts 1975–2005**



**Source:** National Accounts United Nations, PWT6.3, own calculations

## 2.3 Consumption Measures of Structural Transformation

Lastly, we turn to the stylized facts of structural transformation when final consumption expenditure shares are used as a measure of economic activity at the sectoral level. We previously offered two main reasons why final consumption expenditure shares may exhibit different patterns than production value added shares: the presence of savings and investment, in particular in the context of open economies, and the fact that final consumption expenditure is a fundamentally distinct concept from value added produced. The goal of the subsection is to establish that these differences between consumption and production based measures do not matter much for agriculture and services, but can have important implications for manufacturing.

Comparable cross-country panel data on consumption expenditure by sector are much less available than such data on either employment or value added shares. We begin by presenting relatively long time series evidence for the US and the UK in Figure 7. The main message from the plots is that for these two countries, production and consumption measures display very similar behavior, both qualitatively and quantitatively. Specifically, nominal consumption shares for agriculture and services are decreasing and increasing over time, respectively, just as they were in the case for nominal value added shares, and the extent of the changes is quite similar too. Moreover, the consumption share for manufacturing displays an inverted-U shape, just as it did in the case for the nominal value added share for manufacturing. Once again, the quantitative features are also similar, with the peak of the curves occurring at similar values of GDP per capita, and the extent of the decrease after the peak also being similar. One difference between consumption shares and value added shares is that the consumption share for manufacturing tends to be a few percentage points higher than the value added share for manufacturing. This occurs because of the fact that the consumption measure implicitly includes distribution services such as retail trade in its measure of manufacturing consumption.

We next consider two data sets on final consumption expenditure by sector: the OECD Consumption Expenditure Data Base and the Benchmark Studies of the International Comparisons Programme, as reported by the Penn World Table. The OECD data have reasonably long time series for several currently rich countries, namely, Australia, Canada, Japan, Korea, and the United States as well as the seven EU countries Austria, Denmark, Finland, France, Italy,

Netherlands, and the United Kingdom. The Benchmark Studies offer relatively large cross sections for the years 1980, 1985, and 1996. We define the sectors for consumption expenditure following the usual conventions; for example, we use food as the category closest to agriculture; for the details see Appendix A. For each data set, we pool the data and plot the nominal consumption expenditure shares of the three sectors against GDP per capita measured in international dollars.

Figure 8 contains the plots for the OECD data and Figure 9 contains the plots for the Penn World Table data. Two patterns are immediate: the final expenditure share for food tends to decrease with the level of development while the final expenditure share for services tends to increase with development. These two patterns are qualitatively similar to the patterns that we have documented by using the production based measures of economic activity at the sectoral level. However, when we examine the plot for manufacturing consumption we now see some differences. Of particular interest is Korea; whereas it exhibits the same inverted u-shape as the other OECD countries for the production value added share of manufacturing, we see that its consumption share of manufacturing is virtually flat during a period of rapid growth.

The data from the PWT for the manufacturing consumption share effectively show a cloud. While this plot is not necessarily inconsistent with an inverted-U shape for each country coupled with level differences across countries, it suggests that differences between production and consumption measures may be a more common feature of the data in the larger sample of countries. We think this is an important issue that merits further work. If the link between consumption and production measures is different for current developing countries than it was for countries that developed earlier, then this may well have implications for the nature of the development path that these countries follow.<sup>16</sup>

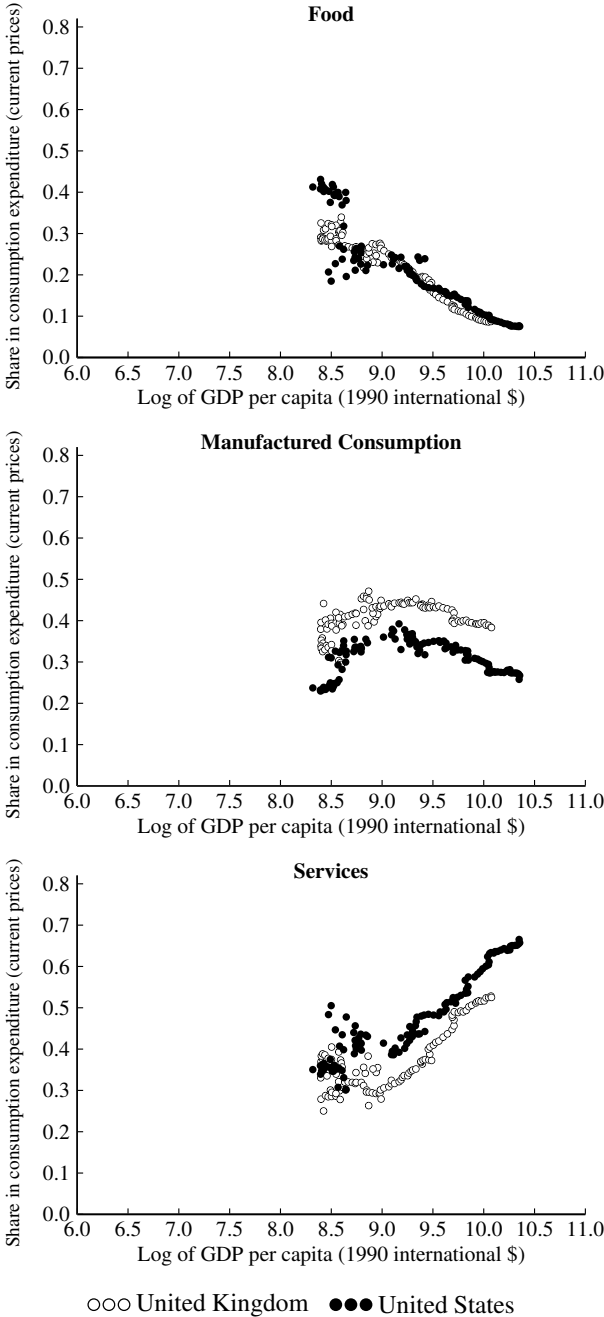
### **3 Modeling Structural Transformation and Growth**

In this section we present a natural extension of the one-sector growth model that incorporates structural transformation. For reasons that will become clear shortly, it turns out to be useful

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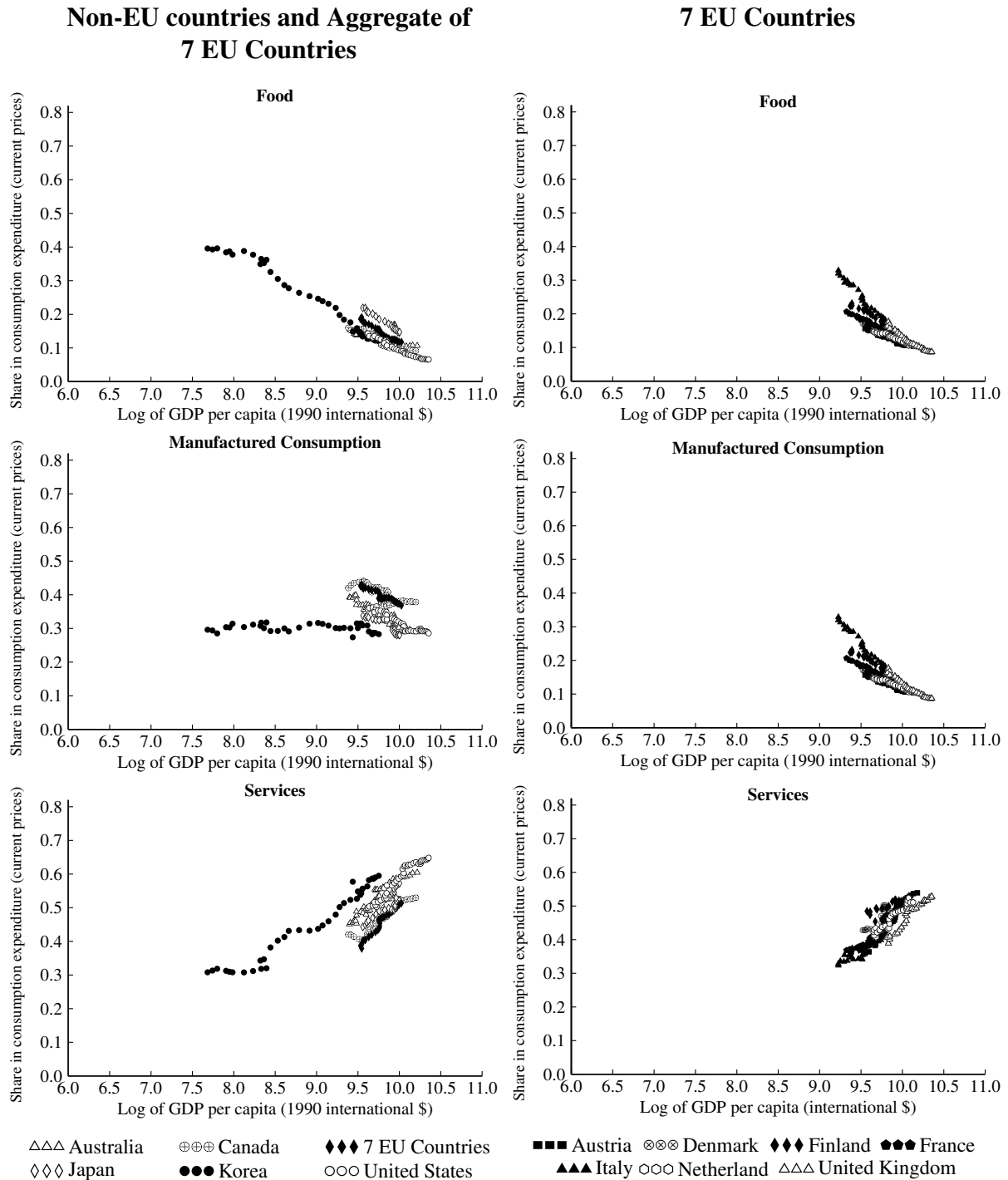
<sup>16</sup>Looking at sectoral employment shares, Bah (2008) documents that the process of structural transformation in many developing countries looks very different than the historical experiences of current rich countries.

**Figure 7: Sectoral Shares of Nominal Consumption Expenditure – US and UK 1900–2008**



**Source:** Various historical statistics, see Appendix A.

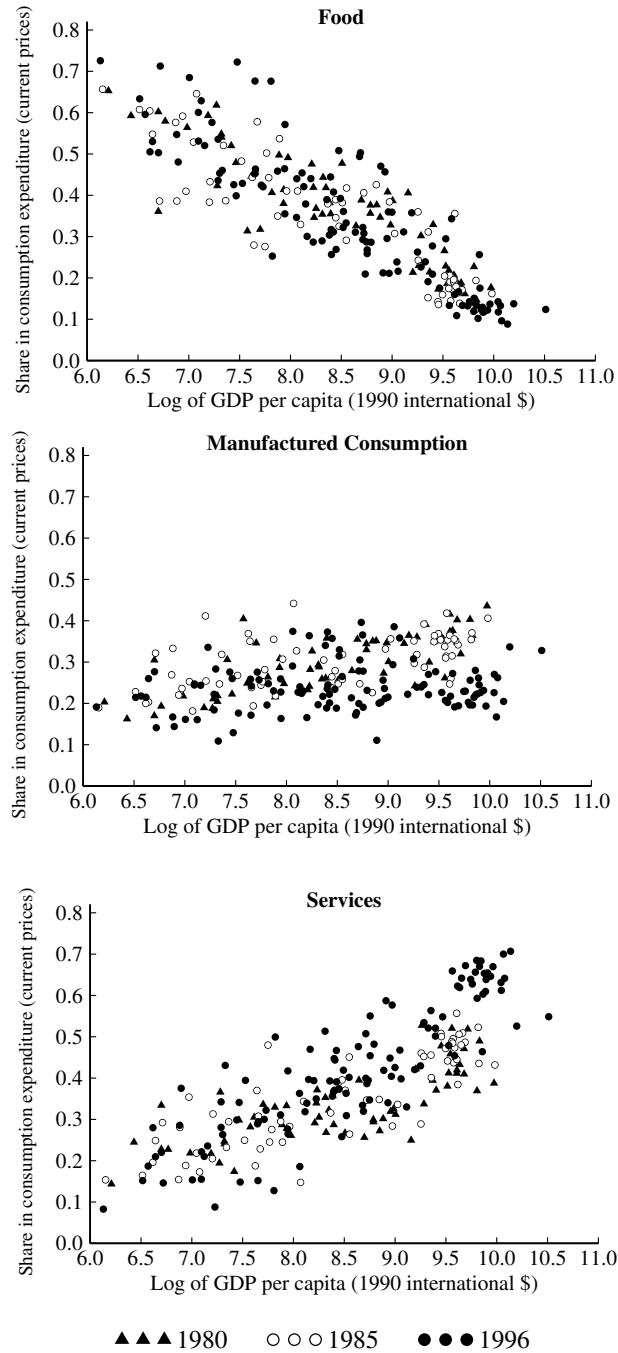
**Figure 8: Sectoral Shares of Nominal Consumption Expenditure – Various Countries, OECD 1970–2007**



**Source:** OECD, EUKLEMS, PWT6.3



**Figure 9: Sectoral Shares of Nominal Consumption Expenditure – Cross Sections from the ICP Benchmark Studies 1980, 1985, 1996**



**Source:** International Comparisons Programme (as reported in PWT)

to first consider a two–sector version of the growth model that has separate consumption and investment sectors. We therefore begin by presenting this model and analyzing some of its properties before we proceed to describe our extension aimed at studying growth and structural transformation.

### 3.1 Background: A Two–Sector Version of the Growth Model

Our presentation of the two-sector growth model closely resembles that in Greenwood et al. (1997). We assume an infinitely lived stand–in household with preferences over consumption sequences  $\{C_t\}$  given by:

$$\sum_{t=0}^{\infty} \beta^t \log C_t \quad (1)$$

where  $0 < \beta < 1$  is the discount factor. Note that, for simplicity, preferences are such that the household does not value leisure. The household is endowed with one unit of productive time and an initial stock of capital,  $K_0$ .

There are two constant–returns–to–scale production functions which describe how consumption ( $C$ ) and investment ( $X$ ) are produced from capital,  $k$ , and labor,  $n$ . It is convenient to follow the literature and impose that the production functions are Cobb–Douglas and have the same capital share:

$$C_t = k_{ct}^{\theta} (A_{ct} n_{ct})^{1-\theta}$$

$$X_t = k_{xt}^{\theta} (A_{xt} n_{xt})^{1-\theta}$$

where  $A_{it}$  represents exogenous labor–augmenting technological progress in sector  $i$ . We adopt the notational convention of using upper–case letters to refer to aggregate variables.

Capital accumulates as usual:

$$K_{t+1} = (1 - \delta)K_t + X_t$$

where  $0 < \delta < 1$  denotes the depreciation rate.

We assume that capital and labor are freely mobile between the two sectors so that feasibility

requires that in each period:

$$K_t = k_{ct} + k_{xt}$$

$$1 = n_{ct} + n_{xt}$$

As is standard, we study the competitive equilibrium for this economy. Although one can obtain the competitive–equilibrium allocations by solving a social planner’s problem, we want to emphasize the role of relative prices and so therefore consider a sequence–of–markets competitive equilibrium in which the price of the investment good is normalized to be equal to one in each period. The price of the consumption good relative to the investment good is denoted by  $P_t$ , the rental rate for capital is denoted by  $R_t$ , and the wage rate is denoted by  $W_t$ . We assume that the household accumulates capital and rents it to firms.

We begin our characterization of the equilibrium by establishing that the capital–to–labor ratios are equalized across sectors at each point in time. To see this note that the first–order conditions for the stand–in firm in sector  $i \in \{c, x\}$  are given by:

$$R_t = P_t \theta \left( \frac{k_{it}}{n_{it}} \right)^{\theta-1} A_{it}^{1-\theta}$$

$$W_t = P_t (1 - \theta) \left( \frac{k_{it}}{n_{it}} \right)^{\theta} A_{it}^{1-\theta}$$

Combining these two equations and rearranging gives an expression for the capital–labor ratio in sector  $i \in \{c, x\}$ :

$$\frac{k_{it}}{n_{it}} = \frac{\theta}{1 - \theta} \frac{W_t}{R_t}$$

Hence, the two capital–to–labor ratios are equal. It follows that the capital–to–labor ratio in each sector is the same as the aggregate capital–to–labor ratio:<sup>17</sup>

$$\frac{k_{ct}}{n_{ct}} = \frac{k_{xt}}{n_{xt}} = K_t \tag{2}$$

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<sup>17</sup>To see this note that

$$\frac{k_{ct}}{n_{ct}} n_{ct} + \frac{k_{xt}}{n_{xt}} n_{xt} = K_t (n_{ct} + n_{xt}) = K_t.$$

Next, we establish that the equilibrium value of the relative price  $P_t$  is pinned down by technology. To see this, divide the first-order conditions for labor from the two sectors by each other and use the fact that sectoral capital-to-labor ratios are equalized. This gives:

$$P_t = \left( \frac{A_{xt}}{A_{ct}} \right)^{1-\theta} \quad (3)$$

Equations (2) and (3) imply that

$$P_t C_t = \left( \frac{k_{ct}}{n_{ct}} \right)^\theta P_t A_{ct}^{1-\theta} n_{ct} = K_t^\theta A_{xt}^{1-\theta} n_{ct}$$

It follows that the model aggregates on the production side, in that we can consider an aggregate production function that produces a single good that can be turned into either consumption or investment via a linear technology with marginal rate of transformation equal to  $P_t$ :

$$Y_t = X_t + P_t C_t = K_t^\theta (A_{xt})^{1-\theta} (n_{xt} + n_{ct}) = K_t^\theta A_{xt}^{1-\theta} \quad (4)$$

Additionally, equation (2) and the first-order conditions for the firm in the investment sector imply that the marginal products of the aggregate production function determine the rental rate of capital and the wage rate:

$$R_t = \theta K_t^{\theta-1} A_{xt}^{1-\theta} \quad (5)$$

$$W_t = (1 - \theta) K_t^\theta A_{xt}^{1-\theta} \quad (6)$$

To characterize the competitive equilibrium further, we turn to the household side. The

household's maximization problem is:<sup>18</sup>

$$\max_{\{C_t, K_{t+1}\}_{t=0}^{\infty}} \sum_{t=0}^{\infty} \beta^t \log C_t \quad \text{s.t.} \quad P_t C_t + K_{t+1} = (1 - \delta + R_t)K_t + W_t$$

Letting  $\mu_t$  denote the current-value Lagrange multiplier on the period  $t$  budget equation, the first-order conditions for  $C_t$  and  $K_t$  are:

$$\begin{aligned} \frac{\beta^t}{C_t} &= \mu_t P_t \\ 1 - \delta + R_t &= \frac{\mu_{t-1}}{\mu_t} \end{aligned}$$

Combining these two equations gives the Euler equation:

$$\frac{1}{\beta} \frac{P_t C_t}{P_{t-1} C_{t-1}} = 1 - \delta + R_t \quad (7)$$

Using equations (4) and (5), equation (7) can be written as a second-order difference equation in the aggregate capital stock  $K_t$ . Given a value for the initial capital stock, this second-order difference equation together with a transversality condition determine the equilibrium sequence of capital stocks.

We are now ready to consider the possibility of a balanced growth path in this model. In order to establish the existence of a balanced growth path we assume that both technologies improve at constant, though not necessarily equal, rates  $\gamma_i > 0$ :

$$\frac{A_{it+1}}{A_{it}} = 1 + \gamma_i, \quad i = c, x$$

The standard definition of balanced growth is that endogenous variables either grow at constant rates or that they remain constant. It turns out that this definition is too strict for models with

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<sup>18</sup>Note that if total consumption grows at a constant rate  $\gamma_c$ , which will be the case along a GBGP, then the household's objective function remains finite, and so is well defined. The reason for this is that

$$\sum_{t=0}^{\infty} \beta^t \log C_t = [\log C_0 + \log(1 + \gamma_c)] \sum_{t=0}^{\infty} \beta^t t < \infty$$

structural transformation because the very nature of structural transformation is that the sectoral composition changes. We therefore follow the literature and use the weaker concept of *generalized balanced growth path* (GBGP), which only requires that the real interest rate is constant. Equation (5) shows that for the real interest rate to be constant,  $K_t$  needs to grow at the same rate as  $A_{xt}$ . If  $K_t$  grows at the constant rate  $\gamma_x$ , then the law of motion for capital implies that  $X_t$  must grow at the same constant rate. Equation (4) then implies that  $P_t C_t$  must also grow at this same rate. Substituting this growth rate into equation (7) pins down the constant value of the rental rate of capital along a GBGP:

$$\frac{1}{\beta}(1 + \gamma_x) = 1 - \delta + R$$

Given a value for  $A_{x0}$ , the condition on the equilibrium rental rate uniquely determines the value of  $K_0$  along a GBGP:

$$K_0 = \left[ \frac{\beta\theta}{(1 + \gamma_x) - \beta(1 - \delta)} \right]^{\frac{1}{1-\theta}} A_{x0} \quad (8)$$

We note several features of this balanced growth path. First, along the GBGP,  $K_t$  and  $C_t$  grow at different rates. In particular, since (3) implies that  $P_t$  grows at gross rate  $[(1 + \gamma_x)/(1 + \gamma_c)]^{1-\theta}$ , and  $P_t C_t$  grows at gross rate  $(1 + \gamma_x)$ , it follows that  $C_t$  grows at gross rate  $(1 + \gamma_x)^\theta (1 + \gamma_c)^{1-\theta}$ , i.e., a weighted average of the two sectoral growth rates in technology. Given that  $X_t$  grows at the same rate as both  $A_{xt}$  and  $K_t$ , it follows that sectoral employment and capital shares are constant along the balanced growth path. In other words, although in this model differential rates of technological progress lead to changes in relative prices of sectoral outputs, these price changes are not associated with any changes in factor allocations over time.

For future reference, it is of interest to note that although we assumed that technological progress in both sectors is constant over time, this is not required for the existence of a GBGP. In fact, because along the GBGP the difference in technological progress only shows up in prices and not in allocations, it follows that the same results would apply even if the growth rate of technological progress in the consumption sector varied over time. This would have no effect on how capital and labor are allocated and would only show up in the behavior of the relative price  $P_t$ . While in this case not all variables would grow at constant rates, it would still

be true that the rental rate on capital would be constant and that  $Y_t$  and  $K_t$  would grow at the same constant rate. Thus, there would still be a GBGP.

## 3.2 A Benchmark Model of Growth and Structural Transformation

We use the model of the previous section as the starting point for our analysis of structural transformation in the context of the growth model.

### 3.2.1 Set up of the Benchmark Model

As in the previous section, we assume an infinitely lived stand-in household that has preferences characterized by (1) and is endowed with one unit of time and an initial capital stock.

Different than in the previous section, we now assume that  $C_t$  is a composite of agricultural consumption ( $c_{at}$ ), manufacturing consumption ( $c_{mt}$ ) and service consumption ( $c_{st}$ ):

$$C_t = \left[ \omega_a^{\frac{1}{\varepsilon}} (c_{at} - \bar{c}_a)^{\frac{\varepsilon-1}{\varepsilon}} + \omega_m^{\frac{1}{\varepsilon}} (c_{mt})^{\frac{\varepsilon-1}{\varepsilon}} + \omega_s^{\frac{1}{\varepsilon}} (c_{st} + \bar{c}_s)^{\frac{\varepsilon-1}{\varepsilon}} \right]^{\frac{\varepsilon}{\varepsilon-1}} \quad (9)$$

where  $\bar{c}_i, \omega_i \geq 0$  and  $\varepsilon > 0$ . The functional form (9) is a parsimonious choice that allows us to capture two features on the demand side that are potentially important for understanding the reallocation of activity across these three sectors: income effects and relative-price effects. In particular, the presence of the two terms  $\bar{c}_a$  and  $\bar{c}_s$  allows for the period utility function to be non-homothetic and therefore the possibility that changes in income will lead to changes in expenditure shares even if prices are constant. The parameter  $\varepsilon$  influences the elasticity of substitution between the three goods and hence the response of nominal expenditure shares to changes in relative prices. Perhaps somewhat restrictively, the above specification imposes that the elasticity of substitution between each pair of goods is the same. Although we could generalize it to accommodate the possibility of different elasticities of substitution, we adopt this simpler specification as our benchmark. We will see below that the simpler specification works fairly well empirically, at least in the context of the post World War II US.

We generalize the previous model to allow for four Cobb–Douglas production functions, one for each of the three consumption goods and one for the investment good. Formally, the

production functions are given by:<sup>19</sup>:

$$c_{it} = k_{it}^\theta (A_{it} n_{it})^{1-\theta}, \quad i \in \{a, m, s\} \quad (10)$$

$$X_t = k_{xt}^\theta (A_{xt} n_{xt})^{1-\theta} \quad (11)$$

There is a tradition in the literature of working with only three production functions, with the assumption that all investment is produced by the manufacturing sector, i.e., that output of the manufacturing sector can be used as either consumption or investment, whereas the output of the other two sectors is only used as consumption. There are two reasons that we have not adopted this specification. First, despite the apparent reasonableness of the claim that investment is to first approximation produced exclusively by the manufacturing sector, it turns out that this is not supported by the data. Moreover, such an assumption is becoming increasingly at odds with the data over time, due at least in part to the fact that software is both a sizeable and increasing component of investment, and most software innovation takes place in the service sector. In fact, total investment has exceeded the size of the entire manufacturing sector in the USA since 2000. The second reason for considering a separate investment sector derives from evidence that technological progress in the investment sector has been more rapid than in the rest of the economy. (See, for example Greenwood et al (1997).) Because the possibility of differential rates of technological progress across sectors will play a key role in the subsequent analysis, we want to allow for the possibility that this rate is different in the investment sector.

Capital is accumulated as usual:

$$K_{t+1} = (1 - \delta)K_t + X_t$$

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<sup>19</sup>We follow much of the literature in abstracting from the differences between physical capital and land and treating land as part of physical capital. We then restrict attention to Cobb–Douglas production functions in capital and labor that have the same capital share in all sectors, which is analytically very convenient, because it implies that we can aggregate the sectoral production functions to an economy–wide Cobb–Douglas production function. In Section 5.1.2 below we will explore to which extent the assumption of equal sectoral capital shares is borne out by the data. For now, we just mention that even if one thinks that sectoral capital shares (where capital includes land) are similar, then there are still important applications for which it is crucial that land is a fixed factor. For such applications one needs to model land and physical capital separately.



As before we assume that capital and labor are freely mobile.<sup>20</sup> With four sectors, the feasibility conditions now take the form:

$$K_t = k_{at} + k_{mt} + k_{st} + k_{xt}$$

$$1 = n_{at} + n_{mt} + n_{st} + n_{xt}$$

### 3.2.2 Equilibrium Properties of the Benchmark Model

We again consider a sequence-of-markets competitive equilibrium in which the price of the investment good is normalized to equal one in each period. The prices of the consumption goods relative to the investment good are denoted by  $p_{it}$ ,  $i \in \{a, m, s\}$ . We again assume that the household accumulates capital and rents it to firms.

Several key properties of the two-sector model that we established above continue to hold in the four-sector model. Specifically, using the same logic as in the previous section, one can show that the capital-to-labor ratios are equalized across the four sectors at each point in time, and are equal to the aggregate capital-to-labor ratio:

$$\frac{k_{it}}{n_{it}} = K_t, \quad i = a, m, s, x \quad (12)$$

Moreover, as before, relative prices are determined by technology:

$$p_{it} = \left( \frac{A_{xt}}{A_{it}} \right)^{1-\theta}, \quad i = a, m, s \quad (13)$$

Using the above results, one can also show that our multi-sector model aggregates on the production side:

$$Y_t = p_{at}c_{at} + p_{mt}c_{mt} + p_{st}c_{st} + X_t = K_t^\theta A_{xt}^{1-\theta} \quad (14)$$

Lastly, the first-order conditions from the firm problems, (6) and (5), are still valid.

On the household side the model is more involved now. In particular, the household problem

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<sup>20</sup>We discuss the case of restricted labor mobility in Section 7.2.

now takes the form:

$$\begin{aligned} & \max_{\{c_{at}, c_{mt}, c_{st}, K_{t+1}\}_{t=0}^{\infty}} \sum_{t=0}^{\infty} \beta^t \log \left[ \omega_a^{\frac{1}{\varepsilon}} (c_{at} - \bar{c}_a)^{\frac{\varepsilon-1}{\varepsilon}} + \omega_m^{\frac{1}{\varepsilon}} (c_{mt})^{\frac{\varepsilon-1}{\varepsilon}} + \omega_s^{\frac{1}{\varepsilon}} (c_{st} + \bar{c}_s)^{\frac{\varepsilon-1}{\varepsilon}} \right]^{\frac{\varepsilon}{\varepsilon-1}} \\ & \text{s.t.} \quad p_{at}c_{at} + p_{mt}c_{mt} + p_{st}c_{st} + K_{t+1} = (1 - \delta + R_t)K_t + W_t \end{aligned}$$

In what follows, we show that this problem can be split into two subproblems: (i) how to allocate total income between total consumption and savings; (ii) how to allocate total consumption expenditure between the three consumption goods. We develop a useful representation in which the first subproblem closely resembles the problem of the household in the two–sector model considered previously.

In order to have a well defined household problem, we need to make sure that the consumption of agricultural goods will exceed the subsistence term  $\bar{c}_a$  in each period. Even if this is the case, a corner solution may still arise in which the household chooses zero consumption of services. For now, we assume that the household problem is well defined and that its solution is interior in all periods. In Proposition 2 below, we offer a formal condition to ensure that this is the case along the GBGP. Essentially this will boil down to requiring that in each period total consumption is “large enough” relative to the two terms  $\bar{c}_a$  and  $\bar{c}_s$ .

The first–order conditions for an interior solution for the three consumption categories are:

$$\frac{1}{C_t} \omega_a^{\frac{1}{\varepsilon}} (c_{at} - \bar{c}_a)^{-\frac{1}{\varepsilon}} C_t^{\frac{1}{\varepsilon}} = \lambda_t p_{at} \quad (15)$$

$$\frac{1}{C_t} \omega_m^{\frac{1}{\varepsilon}} (c_{mt})^{-\frac{1}{\varepsilon}} C_t^{\frac{1}{\varepsilon}} = \lambda_t p_{mt} \quad (16)$$

$$\frac{1}{C_t} \omega_s^{\frac{1}{\varepsilon}} (c_{st} + \bar{c}_s)^{-\frac{1}{\varepsilon}} C_t^{\frac{1}{\varepsilon}} = \lambda_t p_{st} \quad (17)$$

where  $\lambda_t$  denotes the current–value Lagrange multiplier on the budget constraint in period  $t$ . If one raises each of the equations (15)–(17) to the power  $1 - \varepsilon$ , adds them, and uses the definition (9) of  $C_t$ , then one obtains:

$$\frac{1}{C_t} = \lambda_t \left[ \omega_a (p_{at})^{1-\varepsilon} + \omega_m (p_{mt})^{1-\varepsilon} + \omega_s (p_{st})^{1-\varepsilon} \right]^{\frac{1}{1-\varepsilon}} \quad (18)$$

Given that  $\lambda_t$  is the marginal value of an additional unit of expenditure in period  $t$ , it follows that the other term on the right-hand side is naturally interpreted as the price of a unit of composite consumption. In view of this we will define the price index  $P_t$  by:

$$P_t \equiv \left[ \omega_a (p_{at})^{1-\varepsilon} + \omega_m (p_{mt})^{1-\varepsilon} + \omega_s (p_{st})^{1-\varepsilon} \right]^{\frac{1}{1-\varepsilon}} \quad (19)$$

If one adds the three first-order conditions (15)–(17) and uses this definition of  $P_t$ , one also obtains:

$$p_{at}c_{at} + p_{mt}c_{mt} + p_{st}c_{st} = P_t C_t + p_{at}\bar{c}_a - p_{st}\bar{c}_s \quad (20)$$

It follows that the household's maximization problem can be broken down into two subproblems:

**(i) Intertemporal Problem.** Allocate total income among the composite consumption good and savings:

$$\max_{\{C_t, K_{t+1}\}_{t=0}^{\infty}} \sum_{t=0}^{\infty} \beta^t \log C_t \quad \text{s.t.} \quad P_t C_t + K_{t+1} = (1 - \delta + r_t)K_t + w_t - p_{at}\bar{c}_a + p_{st}\bar{c}_s$$

**(ii) Static Problem.** Allocate the period- $t$  consumption expenditure  $P_t C_t$  among the three consumption goods:

$$\max_{c_{at}, c_{mt}, c_{st}} \left[ \omega_a^{\frac{1}{\varepsilon}} (c_{at} - \bar{c}_a)^{\frac{\varepsilon-1}{\varepsilon}} + \omega_m^{\frac{1}{\varepsilon}} (c_{mt})^{\frac{\varepsilon-1}{\varepsilon}} + \omega_s^{\frac{1}{\varepsilon}} (c_{st} + \bar{c}_s)^{\frac{\varepsilon-1}{\varepsilon}} \right]^{\frac{\varepsilon}{\varepsilon-1}}$$

$$\text{s.t.} \quad p_{at}c_{at} + p_{mt}c_{mt} + p_{st}c_{st} = P_t C_t + p_{at}\bar{c}_a - p_{st}\bar{c}_s$$

This representation nicely separates out the growth component of the model from the structural transformation component of the model. From the perspective of balanced growth in the aggregates  $K_t$  and  $C_t$ , the representation looks like the two-sector growth model with the exception of one detail: this economy behaves as if there is a time varying endowment, reflected by the term  $-p_{at}\bar{c}_a + p_{st}\bar{c}_s$ . If this endowment happens to be zero at all dates, then the equivalence to a standard two-sector model is exact. Be that as it may, the Euler equation is still of the form (7). Moreover, although the expression for the relative price  $P_t$  is somewhat more complicated

in the current setting compared to the two–sector model, the equilibrium value of this relative price can still be determined directly from primitives without solving for the full equilibrium.

From the perspective of structural transformation, the above representation implies that we can focus on the solution to the static problem of allocating each period’s consumption expenditure between the three consumption goods. The first–order conditions (15)–(17) characterize the solution to this static problem. For future reference, we note two useful implications of the first–order conditions. First, they impose conditions on the ratios of any two consumption goods:

$$\left(\frac{p_{at}}{p_{mt}}\right)^\varepsilon \frac{c_{at} - \bar{c}_a}{c_{mt}} = \frac{\omega_a}{\omega_m} \quad (21)$$

$$\left(\frac{p_{st}}{p_{mt}}\right)^\varepsilon \frac{c_{st} + \bar{c}_s}{c_{mt}} = \frac{\omega_s}{\omega_m} \quad (22)$$

Second, they impose a condition on the ratio of the expenditure on composite consumption and the expenditure on manufactured consumption:

$$\frac{P_t C_t}{p_{mt} c_{mt}} = \left[ \frac{\omega_a}{\omega_m} \left(\frac{A_{mt}}{A_{at}}\right)^{(1-\theta)(1-\varepsilon)} + 1 + \frac{\omega_s}{\omega_m} \left(\frac{A_{mt}}{A_{st}}\right)^{(1-\theta)(1-\varepsilon)} \right] \quad (23)$$

Equations (21)–(23) will play a key role below when we study the details of structural transformation within the framework of our four–sector model.

### 3.3 Connecting the Benchmark Model to Measures of Structural Transformation

Since we will eventually ask whether versions of this model can help us understand the stylized facts of structural transformation that we documented in Section 2, it is relevant to discuss how to connect the model to the measures from the data that we have previously examined. While this might appear obvious, several subtle issues arise that have to do with the distinctions between consumption and investment and between final expenditure and value added.

When we documented the sectoral patterns of value added in Section 2, we disaggregated total value added into the value added of agriculture, manufacturing, and services and measured

the shares of these three sectors in total value added. If we want to connect our model to these measures of sectoral activity, then it is natural to assume that the sectoral production functions that we have specified in the benchmark model represent value added production functions. But this immediately raises the issue of how to divide the value added of the investment sector between agriculture, manufacturing, and services. The literature often assumes that the entire value added of the investment sector belongs to manufacturing, but we will document below that this assumption is inconsistent with the data. A simple way to see this is by recognizing that in recent years in the United States the value added of the investment sector has exceeded the value added of the manufacturing sector.

Assuming that the sector production functions represent value added production functions leads to a difficulty in trying to connect the model with data on consumption expenditure shares. Because equilibrium requires that  $c_{it} = k_{it}^{\theta}(A_{it}n_{it})^{1-\theta}$ , it would seem natural to identify  $p_{it}c_{it} / \sum_j p_{jt}c_{jt}$  as the model's measure of the nominal consumption share of sector  $i$  in period  $t$ . However, this share is not an appropriate measure for the nominal consumption expenditure share of sector  $i$  as measured in the data. To see why, let us return to the example discussed earlier of the purchase of a cotton shirt. To measure the contribution of this shirt to manufactured final consumption expenditure, we need to aggregate all value added that goes into the production of the shirt through the use of intermediate inputs from each of the three sectors. This requires us to take into account the input–output relationships about how value added is aggregated into final consumption expenditure. In contrast, the above definition of consumption shares includes only the value added that came from the manufacturing sector itself, and so it does not reflect how final consumption expenditure is measured in a world in which each sector uses intermediate inputs from the other sectors.

Alternatively, we could assume that  $p_{it}c_{it} / \sum_j p_{jt}c_{jt}$  in the model does correspond to the nominal consumption expenditure share of sector  $i$  in period  $t$  as measured in the data. But since in equilibrium  $c_{it} = k_{it}^{\theta}(A_{it}n_{it})^{1-\theta}$ , it would then follow that  $p_{it}k_{it}^{\theta}(A_{it}n_{it})^{1-\theta}$  is not an appropriate measure of value added from sector  $i$  in period  $t$  as measured in the data. Returning to the shirt example, this piece of  $c_{mt}$  now reflects the value added components from each of the three sectors that went into producing the final product, and so it cannot be the value added

from one particular sector. In order to maintain consistency, it must be that the production functions summarize the labor and capital from the various stages of production that are used to produce final consumption expenditure. In order to obtain value added shares one would have to use (inverse) input–output relationships to unbundle the final consumption expenditure into its value added components. Moreover, since  $n_{it}$  now reflects all of the labor that went into producing the shirt at each of the various stages of production, it is no longer the case either that  $n_{it}$  is an appropriate measure of the employment share of sector  $i$  in period  $t$ .

The bottom line from this discussion is that if one wants to have a model that can simultaneously address the shares of sectoral employment, value added, and consumption expenditure, then one will need to explicitly include the details of the input–output structure involved in transforming sectoral value added into sectoral consumption expenditure. We have chosen not to do this in order to preserve a greater degree of transparency in the presentation. In view of this, we need to keep in mind that when we discuss the model implications for the measures of structural transformation, we can either connect the production measures (employment shares and value added shares) to the data, implying that the consumption measure (consumption expenditure shares) does not have a close empirical counterpart, or we can connect the consumption measure to the data implying that the two production measures do not have close empirical counterparts. Whichever way we choose, our model will not be able to make statements about all three measures of structural transformation at the same time.

## **4 The Economic Forces Behind Structural Transformation: Theoretical Analysis**

The so–called Kaldor facts regarding balanced growth over long periods of time have lead the profession to focus on specifications of the one–sector neoclassical growth model that generate balanced growth. The evidence that we presented in Section 2 implies that the continuing process of reallocation of activity across sectors coexists with the stable behavior of aggregate variables that characterizes balanced growth. It is therefore perhaps not surprising that the theoretical literature on structural transformation has looked for specifications of the previous

model that give rise to a generalized balanced growth path along which structural transformation occurs. We begin this section by summarizing the results of this theoretical literature and its predictions for the nature of structural transformation. We close this section with a discussion of whether the focus on specifications that deliver exact balanced growth might be too stringent. Independently of this, we believe that the search for specifications that deliver balanced growth and structural transformation has proven useful in helping researchers isolate various forces that are potentially important in shaping structural transformation.

## 4.1 Two Special Cases with Analytical Solutions

Our previous derivations put us in position to easily summarize recent findings in the literature about the joint possibility of generalized balanced growth and structural transformation. In this subsection we focus on two recent papers that emphasize different economic forces behind structural transformation.

### 4.1.1 Preliminaries

If we are to look for a balanced growth path it is natural to limit ourselves to situations in which technological progress is constant. We therefore assume:

$$\frac{A_{it+1}}{A_{it}} = 1 + \gamma_i, \quad i = a, m, c, x \quad (24)$$

As previously noted, even if all aggregates grow at constant rates, it will typically not be the case that all sector level variables grow at constant rates. We therefore follow the literature and focus on generalized balanced growth paths (GBGP), which are defined to be equilibrium paths along which the rental rate of capital is constant, i.e.,  $R_t = R$ . The next result shows that along a GBGP of our model the so-called ‘‘Kaldor facts’’ will hold. These are:  $K_t$  grows at a constant rate;  $Y_t$  grows at a constant rate;  $K_t/Y_t$  is constant;  $R_t$  is constant;  $R_t K_t/Y_t$  is constant.

**Proposition 1.** *If a GBGP exists, then the Kaldor facts hold along the GBGP.*

*Proof.* Since  $R_t$  is constant along a GBGP, it suffices to show that  $K_t$ ,  $Y_t$  and  $X_t$  all grow at rate  $\gamma_x$ .

The fact that  $R$  is constant and equation (5) holds in period  $t$  and  $t + 1$  implies:

$$\frac{A_{xt+1}}{A_{xt}} = \frac{K_{t+1}}{K_t} \quad (25)$$

It follows that  $K_t$  also grows at the constant rate of  $\gamma_x$ . Using  $Y_t = A_{xt}^{1-\theta} K_t^\theta$  we have:

$$\frac{Y_{t+1}}{Y_t} = \left( \frac{A_{xt+1}}{A_{xt}} \right)^{1-\theta} \left( \frac{K_{t+1}}{K_t} \right)^\theta \quad (26)$$

Using equation (25) this gives:

$$\frac{Y_{t+1}}{Y_t} = (1 + \gamma_x)^\theta (1 + \gamma_x)^{1-\theta} = 1 + \gamma_x \quad (27)$$

Constant growth of  $K$  necessarily implies constant growth of  $X$ . The fact that the aggregate technology is Cobb–Douglas implies that factor shares are constant even off a GBGP.  $\square$

We next turn to the issue of whether there are specifications of the model for which a GBGP exists along which structural transformation occurs. At this stage we will simply pose this question from a qualitative perspective. Specifically, we will say that a GBGP exhibits structural transformation if either sectoral employment shares ( $n_{it}$ ) or sectoral value added (or consumption expenditure) shares ( $p_{it}c_{it}/Y_t$ ) are not constant for all three consumption sectors. The issue of generating the “right” patterns of structural transformation, both qualitatively and quantitatively, will be taken up later.

As a starting point it is useful to examine two special cases. The first special case makes the extreme assumption that the three consumption goods are perfect substitutes:  $\bar{c}_a = \bar{c}_s = 0$ ,  $\omega_a = \omega_m = \omega_s$ ,  $A_{at} = A_{mt} = A_{st}$ , and  $\varepsilon \rightarrow \infty$ . In this case the model is identical at the aggregate level to the two–sector model in the previous section, and so it has a unique balanced growth path in terms of  $C_t$  and  $K_t$ . However, since the three consumption goods are perfect substitutes and have identical production functions, the allocation of labor and capital between the three sectors is indeterminate, beyond the restriction that capital–to–labor–ratios must be the same in all sectors with positive output. Because of this indeterminacy it is obviously the case that one can accommodate whatever patterns one desires in terms of changes in either labor allocations



or value added shares across sectors. However, since, as we have seen in Section 2 above, the features of structural transformation appear to be stable over time and across countries, this does not seem a very appealing way to account for structural transformation.

The second special case of interest assumes that  $\bar{c}_a = \bar{c}_s = 0$  and  $\varepsilon = 1$ , so that the preference aggregator is Cobb–Douglas. We do not present the details here, but one can show that the unique balanced growth path has constant sectoral labor and value-added shares. This happens despite the fact that we have not restricted the relative rates of productivity growth among the three consumption sectors. Intuitively, with Cobb–Douglas preferences, employment and value added shares are independent of relative productivities. With sectoral employment and capital shares fixed, differences in relative productivities generate differences in relative outputs, but these differences in output are perfectly offset in terms of value added shares by changes in relative prices. While this special case gives rise to balanced growth and avoids the indeterminacy of the previous case, it does not give rise to structural transformation along the balanced growth path.

In what follows we describe two scenarios that can generate structural transformation along a GBGP. Each of them can be understood as a departure from this second special case.

#### 4.1.2 Case 1: Income Effects and Structural Transformation

Case 1 corresponds to the analysis found in Kongsamut et al. (2001) and represents the extreme scenario in which all structural change is driven by *income effects* that are generated by the terms  $\bar{c}_a$  and  $\bar{c}_s$ . For this case we assume that technological progress is uniform across all consumption sectors ( $\gamma_i = \gamma_j$  for all  $i, j = a, m, s$ ) and that the parameter governing the elasticity of substitution among consumption goods is unity ( $\varepsilon = 1$ ).<sup>21</sup> The consumption aggregator (9) then takes the well-known Stone–Geary form:

$$C_t = \omega_a \log(c_{at} - \bar{c}_a) + \omega_m \log(c_{mt}) + \omega_s \log(c_{st} + \bar{c}_s) \quad (28)$$

With  $\bar{c}_a$  and  $\bar{c}_s$  positive it is easy to see intuitively how one may get structural transformation along a GBGP; as income grows, the non-homotheticity of the demands for the different con-

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<sup>21</sup>Note that  $\varepsilon$  equals the elasticity of substitution only if  $\bar{c}_a = \bar{c}_s = 0$ .

sumption goods will lead to changes in the value-added shares. However, there is a potential issue in obtaining GBG when  $\bar{c}_a$  and  $\bar{c}_s$  are positive. To see this, recall the Euler equation (7) for the household problem. From this equation, if  $R_t$  is constant over time, then it must be that  $P_t C_t$  grows at a constant rate. From the period-budget equation, (20), and noting that factor payments are equal to output, we have:

$$P_t C_t + p_{at} \bar{c}_a - p_{st} \bar{c}_s = K_t^\theta A_{xt}^{1-\theta} + (1 - \delta)K_t - K_{t+1} \quad (29)$$

Since the right-hand side grows at rate  $\gamma_x$ ,  $P_t C_t + p_{at} \bar{c}_a - p_{st} \bar{c}_s$  must also grow at rate  $\gamma_x$ . If  $p_{a0} \bar{c}_a - p_{s0} \bar{c}_s$  is not zero, then  $p_{at} \bar{c}_a - p_{st} \bar{c}_s$  will grow at rate  $\gamma_x$  only if relative prices also grow at rate  $\gamma_x$ . However, this contradicts the fact that  $p_{at}$  and  $p_{st}$  both grow at gross rate  $[(1 + \gamma_{xt})/(1 + \gamma_{ct})]^{1-\theta}$ , which is implied by expression (13). Hence, balanced growth requires that  $p_{a0} \bar{c}_a - p_{s0} \bar{c}_s = 0$ , which is equivalent to:

$$\frac{\bar{c}_a}{\bar{c}_s} = \left( \frac{A_{a0}}{A_{s0}} \right)^{1-\theta} \quad (30)$$

Note that since both relative prices grow at the same rate, this condition implies that  $p_{at} \bar{c}_a - p_{st} \bar{c}_s = 0$  at all dates  $t$ .<sup>22</sup>

Given condition (30), equation (29) simply requires that  $P_t C_t$  grows at rate  $\gamma_x$ . From the perspective of balanced growth this economy then looks very much like the two-sector model that we considered in the previous section. In particular, similar to that two-sector model, the share of labor and capital devoted to consumption versus investment is constant along a GBGP.

We make two remarks regarding condition (30). First, note that if either of  $\bar{c}_a$  or  $\bar{c}_s$  is positive, then they must both be positive. As we discuss in a later section, many papers have implicitly assumed that  $\bar{c}_a > 0$  and  $\bar{c}_s = 0$ , which is inconsistent with condition (30). Second, this condition relates the parameters of preferences and technology to each other, and is therefore somewhat of a ‘‘fragile’’ condition. We shall return to this point later in this section.

Next we consider whether structural transformation occurs along the GBGP. To examine

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<sup>22</sup>This point illustrates that the assumption of the same rate of technological progress in the agriculture and service sectors is a necessary condition and not merely a simplification.

this note that if  $\varepsilon = 1$ , then (21)–(22) imply the Stone–Geary demand system:

$$c_{at} = \omega_a \frac{P_t C_t}{P_{at}} + \bar{c}_a \quad (31)$$

$$c_{mt} = \omega_m \frac{P_t C_t}{P_{mt}} \quad (32)$$

$$c_{st} = \omega_s \frac{P_t C_t}{P_{st}} - \bar{c}_s \quad (33)$$

Moreover, the assumption that all consumption sectors grow at the same rate implies that the relative prices of the three consumption goods are constant:

$$\frac{p_{it}}{P_t} = \frac{p_{i0}}{P_0}, \quad i \in \{a, m, s\}$$

Hence  $c_{at}$ ,  $c_{mt}$ , and  $c_{st}$  grow at a slower rate, at the same rate, and at a faster rate than  $C_t$ , respectively. Given that the relative prices of the three consumption goods are constant, it follows that  $p_{it}c_{it}/P_t C_t$  is decreasing for agriculture, constant for manufacturing and increasing for services. Since total consumption expenditures are a constant share of total output, it follows that these properties also carry over to both  $n_{it}$  and  $p_{it}c_{it}/Y_t$ .

In summary, and more formally, we have the following result:

**Proposition 2.** *Assume that condition (30) holds and that*

$$\bar{c}_s \leq \omega_s \left( \frac{A_{s0}}{A_{x0}} \right)^{1-\theta} \left[ K_0^\theta A_{x0}^{1-\theta} - (\gamma_x + \delta)K_0 \right] \quad (34)$$

where  $K_0$  is given by (8).

*Then there is a unique GBGP. Along the GBGP, the employment and nominal value added shares of the investment sector are constant. The employment and nominal value added shares are decreasing for agriculture, constant for manufacturing and increasing for services.*

*Proof.* We start by noting that it is straightforward to show that (8) implies that  $K_0^\theta A_{x0}^{1-\theta} > (\gamma_x + \delta)K_0$ . Hence,  $P_0 C_0 = K_0^\theta A_{x0}^{1-\theta} - (\gamma_x + \delta)K_0 > 0$  and condition (34) is well defined. Condition (34) ensures that the right-hand side of (33) is positive at  $t = 0$ . Since the economy grows while relative prices remain constant, this implies that the right-hand side is positive for

all  $t$ . In this case, equations (31)–(33) are well defined and they have a unique interior solution for  $c_{at}, c_{mt}, c_{st}$ . The existence of a unique GBGP and the statements about the shares then follow directly from the previous discussion.  $\square$

### 4.1.3 Case 2: Relative Price Effects and Structural Transformation

The second case that we consider corresponds to the analysis found in Ngai and Pissarides (2007).<sup>23</sup> Whereas the previous case generated structural transformation purely via income effects and asked whether this could be consistent with balanced growth, Ngai and Pissarides consider the polar extreme case in which structural transformation is generated purely from relative price effects and ask whether this can be consistent with balanced growth. Accordingly, they assume that  $\bar{c}_a = \bar{c}_s = 0$ . In order to have relative price effects operating it is clearly necessary to have differential rates of technological progress among the three consumption goods sectors, so no restrictions will be placed on the relative values of  $\gamma_i$ . Given our earlier discussion, however, we know that  $\varepsilon$  will have to take on a value other than unity.

The analysis of this case follows directly from our analysis of the two–sector model. Specifically, if the values of  $\gamma_a, \gamma_m,$  and  $\gamma_s$  are different, then the price index  $P_t$  will not grow at a constant rate. However, as noted at the end of the section on the two–sector model, this has no bearing on the existence of a unique GBGP; there still is a unique GBGP that features a constant share of labor and capital allocated to total consumption. Along the GBGP the value of  $P_t C_t$  will grow at the constant rate  $\gamma_x$  even though neither component grows at a constant rate.

To assess the implications for structural transformation we again turn to equations (21) and (22). Using equation (13) for relative prices, these two equations can now be written as:

$$\frac{c_{at}}{c_{mt}} = \frac{\omega_a}{\omega_m} \left( \frac{A_{at}}{A_{mt}} \right)^{\varepsilon(1-\theta)} \quad (35)$$

$$\frac{c_{st}}{c_{mt}} = \frac{\omega_s}{\omega_m} \left( \frac{A_{st}}{A_{mt}} \right)^{\varepsilon(1-\theta)} \quad (36)$$

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<sup>23</sup>This work builds on the important earlier contribution of Baumol (1967).

Noting that  $c_{it} = K_t^\theta A_{it}^{1-\theta} n_{it}$ , we also have:

$$\frac{n_{at}}{n_{mt}} = \frac{\omega_a}{\omega_m} \left( \frac{A_{mt}}{A_{at}} \right)^{(1-\varepsilon)(1-\theta)} \quad (37)$$

$$\frac{n_{st}}{n_{mt}} = \frac{\omega_s}{\omega_m} \left( \frac{A_{mt}}{A_{st}} \right)^{(1-\varepsilon)(1-\theta)} \quad (38)$$

Recalling that labor allocated to the overall consumption sector is constant, it follows that if  $\varepsilon = 1$  we have the earlier result that the  $n_{it}$  are constant in each of the three consumption sectors. So too are the values of  $p_{it}c_{it}/P_t C_t$  and  $p_{it}c_{it}/Y_t$ . If  $\varepsilon$  differs from one, then the model can generate structural transformation along a GBGP as long as the rates of technological progress differ among the three consumption sectors. In contrast to Case 1, it is not true in this case that  $c_{mt}$  is a constant proportion of  $C_t$ , nor is true that  $C_t$  grows at a constant rate. Without imposing some additional structure one cannot say more about the nature of structural transformation that occurs.

To simplify exposition, we focus on the special case in which technological progress is strongest in agriculture and weakest in services, that is,  $\gamma_a > \gamma_m > \gamma_s$ . If in addition we assume that  $\varepsilon < 1$ , then the above expressions imply that along a GBGP the values of  $n_{it}$ ,  $p_{it}c_{it}/P_t C_t$  and  $p_{it}c_{it}/Y_t$  are decreasing for agriculture and increasing for services. The behavior of these values for manufacturing is ambiguous in terms of the direction of change, but the size of the change is bounded by the sizes of the change in the other two sectors.

More formally, we summarize the above discussion with the following proposition.

**Proposition 3.** *Let  $\bar{c}_a = \bar{c}_s = 0$ ,  $\varepsilon < 1$ ,  $\gamma_a > \gamma_m > \gamma_s > 0$ , and  $\gamma_x > 0$ .*

*There is a unique GBGP. Along the GBGP, the shares of employment and nominal value added (in current prices) of the investment sector are constant; the shares of employment and nominal value added (in current prices) of the consumption sectors behave as follows: the agricultural shares decline; the services shares rise; the manufacturing shares decrease less than the agricultural shares and increase less than the service shares.*

#### 4.1.4 Qualitative Assessment

The previous subsections outlined two different theories of structural transformation in the context of generalized balanced growth. Although we postpone a more rigorous assessment of the economic mechanisms implicit in these two theories until a later section, it is still of interest at this point to assess the extent to which each of the theories can account for some of the broad patterns that we documented in Section 2. This assessment comes out fairly bleak, in that we will conclude that the two leading models of structural transformation that we have reviewed so far have serious trouble accounting for the basic regularities of structural transformation.

Given the qualifications that we have noted previously in connecting the model with data, we keep in mind that we can either connect the production measures (employment shares and value added shares) to the data, implying that the consumption measure (consumption expenditure shares) does not have a close empirical counterpart, or we can connect the consumption measure to the data, implying that the two production measures do not have close empirical counterparts. Whichever way we choose to proceed, our benchmark model will not be able to make statements about all three measures of structural transformation at the same time.

We begin with the model of Kongsamut et al. (2001). Since the investment sector uses a constant share of labor and accounts for a constant share of (nominal) output, it will not influence the trend behavior of any quantities if it is allocated across the three sectors in constant proportions. Assuming this and starting with the nominal production measures, we conclude that the model can account for the increase in the service sector shares and the decrease in the agricultural sector measures along its GBGP, but it does not generate an inverted U shape for the manufacturing sector measures. If one allows for the investment share of manufacturing to decrease over time, as is true in the US data, then the model could generate a decline in both production measures for manufacturing. The increasing share of services in investment would only accentuate the rising employment and nominal value added shares for services. Turning to the nominal consumption expenditure measures, the model can account for the increase in the service share, the near constancy of the manufacturing share, and the decrease in the agricultural share.

The model of Kongsamut et al. (2001) has two additional implications that are counterfac-

tual. First, along its generalized balanced growth relative prices need to be constant. It follows that along a GBGP the real measures of structural transformation must display exactly the same properties as the nominal measures, which means that the model cannot account for the quantitative differences between the nominal and the real measures. Second, the model of Kongsamut et al. (2001) implies that in sufficiently poor economies, the household will consume a zero quantity of services and employment in services will also be zero. In contrast, we saw in Section 2 that even in the poorest countries service employment and value added are bounded away from zero.

Next we turn to the model of Ngai and Pissarides (2007). Once again we note that since along the GBGP the share of labor devoted to investment is constant and the nominal share of investment in output is constant, any constant allocation of investment across the three sectors will not influence any of the trend properties. In this case, given the previously assumed ranking for the rates of technological progress, we conclude that structural transformation along the model's GBGP is qualitatively consistent with the evidence for employment and nominal value added shares in both agriculture and services, but not necessarily for manufacturing. The reason for this is that the model does not have definitive predictions regarding the pattern for manufacturing: while it is possible that it generates an inverted-U shape for the manufacturing shares of employment and nominal valued added, this need not be the case. Turning to the nominal consumption expenditure measures, the model can account for the increase in the service share and the decrease in the agricultural share, but again it does not have definitive predictions regarding the nominal share of manufactured consumption goods.

In contrast, the model of Ngai and Pissarides (2007) does not perform well at all in accounting for the behavior of the real shares, irrespective of whether we use production or consumption related measures. The reason for this is that the assumption of an inelastic CES utility function – i.e.,  $\varepsilon \in [0, 1)$  – implies that nominal and real value added shares necessarily move in opposing directions. To see this, consider for example the effects of an increase in the productivity in manufacturing relative to the productivity in services, which leads to a decrease in the price of manufacturing relative to services. The assumption of a CES utility function implies that this decrease in the relative price cannot lead to a decrease in the consumed quantities of manufac-

tured goods relative services. The reason is that if  $\varepsilon > 0$  then the consumed relative quantity of manufactured goods increases; if  $\varepsilon = 0$  (which is the extreme case of Leontief preferences) then it remains the same. Turning now to the nominal value added share of manufactured goods, since Ngai and Pissarides assumed that demand is inelastic, the increase in the relative quantity of manufacturing does not compensate for the decrease in the relative price, and so the share of nominal value added from manufacturing falls. As a result, given the assumptions about relative TFPs and relative prices, the model cannot generate a decrease in the real quantity of manufacturing (or agriculture) relative to services.

In summary, although each of these two specifications can account for some of the qualitative patterns that we documented previously, neither of them is able to match all of the patterns.

## 4.2 Alternative Specifications

In the preceding analysis, we have summarized the results from two papers regarding the possibility of simultaneously having structural transformation and generalized balanced growth. We chose these two papers because they illustrate two different channels through which expenditure shares may change over time: income effects and relative-price effects. In this subsection we describe some alternative formulations of these two channels that have appeared in the literature.

### 4.2.1 Other Specifications Emphasizing Income Effects

Above we chose a specification of preferences where income effects on expenditure shares were captured by the two terms  $\bar{c}_a$  and  $\bar{c}_s$ . While we think that this is a tractable and transparent way of introducing income effects, there are several alternative specifications of nonhomothetic preferences in the literature that can generate income effects. Here we discuss some examples.

In the first quantitative analysis of structural transformation within the framework of the growth model, Echevarria (1997) generated income effects by using the following alternative specification of the intertemporal utility function:

$$\sum_{t=0}^{\infty} \beta^t [\alpha_a \log c_a + \alpha_m \log c_m + \alpha_s \log c_s - \eta (c_a^{-\rho_a} + c_m^{-\rho_m} + c_s^{-\rho_s})]$$



where  $\alpha_i > 0, \eta, \rho_i \geq 0$ . If  $\eta = 0$  then the preferences reduce to a Cobb–Douglas specification, but if  $\eta > 0$  and at least one of the  $\rho_i > 0$  then the preferences are not homothetic. To see some of the features of this specification it is useful to examine the properties of the marginal utility of good  $i$ , which is given by:

$$MU_i(c_i) = \alpha_i c_i^{-1} + \eta \rho_i c_i^{-1-\rho_i} \quad (39)$$

Note first that the marginal utility of each good will be infinite for zero consumption quantities, implying that the household chooses interior consumption quantities. The second term is positive if  $\eta \rho_i > 0$ . In this case, it goes to infinity as  $c_i$  becomes arbitrarily small and it goes to zero as  $c_i$  becomes arbitrarily large.

If, as in Echevarria’s calibration,  $\eta > 0$  and  $\rho_a > \rho_m > \rho_s = 0$ , then at low levels of income (and hence of consumption), there is a force in favor of higher  $c_a$  and  $c_m$  and of lower  $c_s$ , and the force is stronger for  $c_a$  than for  $c_m$ . In contrast, at high levels of income this force disappears. Intuitively, one can use the parameters  $\eta$  and  $\rho_i$  to achieve the same qualitative effects that are generated by the parameters  $\bar{c}_a$  and  $\bar{c}_s$  in our benchmark model.

The main advantage of Echevarria’s specification of period utility is that an interior solution to the static period problem exists for any positive level of income. This is in contrast to what happens in our benchmark model, since if  $\bar{c}_a > 0$  and the present value of income is lower than the present value of  $\{p_{at}\bar{c}_a\}$ , then the household cannot afford to purchase at least  $\bar{c}_a$  units of the agricultural good in all periods and our period utility will not be defined in at least one period. From an analytical perspective, the disadvantage of Echevarria’s specification is that it is not consistent with generalized balanced growth. The reason for this is the presence of the term  $\eta c_j^{-\rho_j}$  in the period utility function. If  $\eta = 0$ , then period utility is of the homothetic log form and a GBGP exists. In contrast, if  $\eta > 0$ , then it is impossible for the value of total consumption,  $\sum_{j \in \{a,m,s\}} p_{jt} c_{jt}$ , to grow at the same constant rate at which technological progress grows. As we saw in Section 3.2 above, this would be required for a GBGP with constant real interest rate to exist.

A recent paper by Boppart (2011) explores more general preferences that are consistent

with balanced growth. In particular, Boppart specifies indirect period utility functions that fall into the class of “price-independent-generalized-linearity” preferences defined by Muellbauer (1975, 1976). These preferences are more general than Gorman preferences in that they generate nonlinear Engel curves. Nonetheless they aggregate and allow for a stand-in household. There are two advantages of using “price-independent-generalized-linearity” preferences in the context of structural transformation. First, they avoid the awkward feature of our benchmark specification that can lead to utility not being defined for sufficiently small income. Second, as Boppart establishes, they are consistent with balanced growth if the technology side is as we specified it above.

A different approach to generating income effects is Foellmi and Zweimüller (2008). Whereas our benchmark model implicitly aggregated individual consumption goods into three broad sectors and defined preferences over the amounts of the three resulting aggregates, these authors specify preferences over an unbounded mass of potential consumption goods. Preferences are such that for each good, marginal utility is finite at zero consumption and decreases to zero at some finite satiation level of consumption. Over time, as income increases, the mass of goods that are consumed increases, so that there is adjustment along both the intensive and the extensive margin. The order in which the goods will be introduced is uniquely determined by the model’s primitives: all of the goods are symmetric from the perspective of production but are given different weights in preferences.<sup>24</sup>

The fact that new goods are consumed over time implies that labor will necessarily be reallocated across activities over time. In terms of basic economic forces, the key mechanism at work comes from the fact that different goods have different income elasticities. Different than in the specification of our benchmark model, however, any particular good in this model will have an income elasticity of zero asymptotically since at some date satiation will be reached.

In order to connect their model to the standard facts of structural transformation, Foellmi and Zweimüller (2008) need to map individual goods into the three broad sectors. If they assume that agricultural goods are disproportionately the goods with high weights, that services are disproportionately the goods with low weights, and that manufacturing goods lie “in be-

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<sup>24</sup>This type of preferences is sometimes called “hierarchical preferences”. It was first used by Murphy et al. (1989).

tween” these two, then they can match the qualitative patterns presented earlier. As income grows and more of the less weighted goods are consumed, one obtains a declining share for agricultural goods, an increasing share for services, and an inverted U shaped pattern for manufacturing. Foellmi and Zweimüller can also generate balanced growth with relatively standard assumptions. Specifically, if they assume that the weighting function on different goods has a power form and there is constant labor augmenting technological progress that is common to the production of all goods, then their model gives rise to a GBGP. As they discuss in their paper, the assumption of a power function for the weighting function is analogous to the assumption of a constant elasticity utility function in the context of the standard one–sector growth model.

Relative to the results that we derived previously about income effects and structural transformation, the specification of Foellmi and Zweimüller (2008) delivers balanced growth and structural transformation in a more robust manner, in the sense that it does not need a condition similar to (30) that imposes a restriction on the parameters of preferences and technology. Moreover, it can also deliver an inverted U shaped relationship between GDP per capita and the manufacturing shares. But a limitation of the specification of Foellmi and Zweimüller (2008) is that modeling structural transformation at the level of individual goods does not provide much guidance for how to connect the model with data at the level of broad sectors.<sup>25</sup>

Hall and Jones (2007) also develop a framework that can give rise to non–homothetic demand functions and income effects, though their focus is specifically on the rise of spending on health care, as opposed to the more general process of structural transformation. Nonetheless this is of interest in the current context since increases in health care account for a significant part of the overall increase in the size of the service sector. In the basic model of Hall and Jones, utility in the current period is derived from a single good that represents all non–health consumption. The period utility function is homothetic and health consumption in period  $t$  provides no direct utility flow in period  $t$  but does influence the probability of survival to next period. Intuitively, this model has features akin to the model with intensive–extensive margins

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<sup>25</sup>Buera and Kaboski (2011) adopt a similar preference structure as Foellmi and Zweimüller (2008) in that they stress the introduction of new goods and adjustment along the extensive margin. Other aspects of their analysis are quite different, however. We discuss their model in more detail in Section 7. For now we simply note that Buera and Kaboski derive an explicit mapping from their preferences to a reduced–form representation of preferences over goods and services. The interesting feature of this mapping is that it includes a term that is analogous to our term  $\bar{c}_s$ , but rather than being a constant, its value changes over time as technological progress occurs.

that we discussed above. Specifically, a household can adjust along the intensive margin by spending more on consumption, or along the extensive margin by spending more on health care and therefore increasing the expected number of periods in which consumption occurs. As the level of consumption increases, the marginal utility from additional consumption at the intensive margin decreases relative to the marginal utility of living an additional period. This can generate an increasing expenditure share for health consumption as incomes rise, and therefore look like a model that features a non-homothetic period utility function over health and non-health consumption.<sup>26</sup>

#### 4.2.2 Other Specifications Emphasizing Relative Price Effects

In this subsection we describe the work of Acemoglu and Guerrieri (2008), who consider a multi-sector version of the growth model in which the sectors have different capital shares but the same rate of exogenous technical change. As capital is accumulated, the relative prices of the more capital intensive goods decline, leading to sustained movements in the relative prices of the different goods. The economics of this model are similar to those in the model of Ngai and Pissarides (2007) except that the underlying cause of the relative price movements is different.

Here we sketch the basic idea within our benchmark model. Assume that technological progress is uniform across the three consumption sectors and define  $A_t$  by  $A_t \equiv A_{it}^{1-\theta_i}$  for  $i \in \{a, m, s\}$ . As just noted, we allow the capital shares to differ across sectors so that the sectoral production functions (10) become:

$$c_{it} = A_t k_{it}^{\theta_i} n_{it}^{1-\theta_i}, \quad i \in \{a, m, s\} \quad (40)$$

All other features of the environment are the same as in the benchmark model described earlier.

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<sup>26</sup>In a recent paper, Lawver (2011) uses a version of the model of the model of Hall and Jones (2007) to measure the increase in the quality of health consumption.

The first-order conditions for the stand-in firm in sector  $i \in \{a, m, s\}$  are now given by:

$$R_t = p_{it}\theta_i A_t \left(\frac{k_{it}}{n_{it}}\right)^{\theta_i-1} \quad (41)$$

$$W_t = p_{it}(1 - \theta_i)A_t \left(\frac{k_{it}}{n_{it}}\right)^{\theta_i} \quad (42)$$

Dividing these equations by each other gives:

$$\frac{1 - \theta_i}{\theta_i} \frac{k_{it}}{n_{it}} = \frac{1 - \theta_j}{\theta_j} \frac{k_{jt}}{n_{jt}} \quad (43)$$

Two implications follow from this equation. First, sectors with larger capital shares have larger capital-labor ratios, and second, the capital-labor ratio grows at the same rate in all sectors.

To derive the implications for relative prices, substitute (43) into (42) and rearrange to yield:

$$\frac{p_{it}}{p_{jt}} = \Omega_{ij} \left(\frac{k_{it}}{n_{it}}\right)^{\theta_j-\theta_i} \quad i, j \in \{a, m, s\} \quad (44)$$

where  $\Omega_{ij}$  is a constant that depends on the capital shares. Since the capital-labor ratios of all sectors grow at the same rate, equation (44) implies that for any pair of sectors, the relative price of the sector with the higher capital share decreases as the aggregate capital stock grows. If one assumes:

$$\theta_a > \theta_m > \theta_s \quad (45)$$

it follows that the price of services relative to manufacturing and of manufacturing relative to agriculture will both increase over time. This implication is of course analogous to what we derived in the context of the Ngai-Pissarides model when we assumed that  $\gamma_a > \gamma_m > \gamma_s$ .

Like the model of Ngai and Pissarides (2007), although this specification can account for the changes in nominal value added shares, it cannot account for the changes in real value added shares. Moreover, it cannot generate the patterns in sectoral employment shares either. To see why, note that using (43), it is straightforward to show that:

$$K = \left( \sum_{j=x,a,m,s} \frac{\theta_j}{1 - \theta_j} n_j \right) \frac{1 - \theta_i}{\theta_i} \frac{k_i}{n_i} \quad (46)$$

Solving this expression for  $k_i/n_i$  and substituting the result into equation (40) gives:

$$c_{it} = A_t K_t^{\theta_i} \left( \frac{\frac{\theta_i}{1-\theta_i}}{\sum \frac{\theta_j}{1-\theta_j} n_j} \right) n_{it}, \quad i \in \{a, m, s\}$$

In the polar case of Leontief utility,  $c_{it}/c_{jt}$  is constant, so the previous equation implies that  $n_{it}/n_{jt}$  is constant too. For positive elasticities of substitution, changes in relative quantities are in the opposite direction of changes in relative prices. In other words, in the model of Acemoglu and Guerrieri there cannot be structural transformation in terms of employment that is consistent with the fact that service employment increased at the same time as which its relative price increased too.

One important additional difference relative to the specification of Ngai and Pissarides (2007) is that the model of Acemoglu and Guerrieri (2008) has exact GBGP only asymptotically, and so the best we can hope for in this model is *approximate* generalized balanced growth. The next subsection discusses the difference between approximate and exact generalized balanced growth in more detail.

### 4.3 Approximate versus Exact Generalized Balanced Growth

Up to this point, our discussion has focused on analytic results concerning the possibility of jointly having generalized balanced growth and structural transformation. This is a natural starting point given the emphasis that the literature using the one-sector growth model places on balanced growth and that conditions under which balanced growth results in the one-sector model are relatively weak – constant returns to scale production with labor augmenting technical change and a period utility function with a constant intertemporal elasticity of substitution. The results that we have presented above for multi-sector models, however, have made it apparent that the conditions for jointly having generalized balanced growth and structural transformation become considerably more stringent – we now need that all production functions are Cobb–Douglas with the same capital share, that the period utility function exhibits a unitary elasticity of substitution, and in some cases that there is a particular relationship between

preference and technology parameters. To the extent that there is good reason to believe that many of these conditions are not satisfied, models that impose them may be missing some key features of reality. In fact, some authors have dismissed income effects as an important source of structural transformation on the grounds that they are consistent with generalized balanced growth only under very fragile cross–restrictions on technology and preferences such as the one imposed in (30).

The previous discussion suggests that it may be ill advised to insist on generalized balanced growth in the context of structural transformation. To the extent that (generalized) balanced growth is merely a good approximation to what we see in the data in various countries over long periods of time, the more relevant question is whether there are specifications that can deliver structural transformation and *approximate* generalized balanced growth, which may occur under much less stringent conditions than exact generalized balanced growth.

To date there has not been much systematic analysis of the extent to which approximate generalized balanced growth is a robust feature of multi–sector versions of the growth model along the lines of those that we have considered. But several cases in the literature suggest that approximate generalized balanced growth may in fact be quite robust. To begin with, Kongsamut et al. (2001) consider numerical examples that depart from the exact conditions needed for generalized balanced growth in their setting and find that the equilibrium path does not deviate much from generalized balanced growth. In a similar context, Gollin et al. (2002) study a two–sector model with subsistence consumption in the agricultural sector but not in the other sector – a clear violation of the conditions needed to generate GBGP, but find relatively small variations of the interest rate when their model is calibrated to match the US data over the post 1950 period. Moreover, although the model in Acemoglu and Guerrieri (2008) only has an asymptotic GBGP, the results that they report for numerical simulations suggest that the model’s behavior along a transition path is not that different from balanced growth.

The models just discussed have the feature that asymptotically structural transformation ceases to occur. That is, in these models structural transformation occurs either along the generalized balanced growth path (in case of Kongsamut et al. (2001)) or along the transitional path (in case of Gollin et al. (2002) and Acemoglu and Guerrieri (2008)) that the economies take in

equilibrium when they converge to the exact balanced growth path. Since we observe (approximate) balanced growth and structural transformation over very long periods in the data, it follows that any model that generates structural transformation purely while it is converging to an exact balanced growth path must have very long lived dynamics in order to capture reality.<sup>27</sup>

## **5 The Economic Forces Behind Structural Transformation: Empirical Analysis**

The previous section has focused on models that could generate (approximate) generalized balanced growth and structural transformation as simultaneous outcomes. The various models that we reviewed emphasize different theories for the reallocation of activity across sectors that accompanies growth. In one class of theories, including Kongsamut et al. (2001), the driving force is uniform technological progress, and the key propagation mechanism comes from income effects. In Ngai and Pissarides (2007), the key driving force is technological progress that differs across sectors and the key propagation mechanism comes from substitution effects. In Acemoglu and Guerrieri (2008), the driving force is again uniform technological progress, and the propagation mechanism is a combination of differential capital intensities in production and substitution effects in consumption.

Rather than focusing narrowly on the conditions required to generate exact balanced growth, we believe that the key to developing quantitative theories of structural transformation is to develop quantitative assessments of the various driving forces and propagation mechanisms that the literature has identified as potentially important. In this section we summarize the recent progress in this effort. We break this section into two subsections. The first subsection considers the direct evidence regarding differences in rates of technological progress and differences in capital shares. The second subsection considers the more general issue of the relative importance of income versus substitution effects.

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<sup>27</sup>Note that this statement does not apply to the model of Ngai and Pissarides (2007) which exhibits structural transformation along the exact balanced growth path also in the limit.



## 5.1 Technological Differences Across Sectors

In this subsection we consider the evidence regarding technological differences across sectors along the two dimensions highlighted by the previous theories: differences in technological progress and differences in capital shares. We also assess the extent to which these differences are appropriate to generate the qualitative features found in the data regarding structural transformation.

### 5.1.1 Sectoral TFP Growth

Assumptions about TFP growth at the sectoral level played an important role in both of the theories of structural transformation that we highlighted. It is therefore of interest to ask what the empirical evidence is regarding relative growth rates in sectoral TFP. Although this would seem to be a relatively straightforward exercise, it is actually challenging to verify the properties of TFP growth in sectoral value added production functions in a cross-country setting. The main reason is that calculating sectoral TFP's requires data on real value added, capital and labor inputs, and the factor shares at the sector level. Unfortunately, these data are unavailable for most countries. One of the many issues is that in order to compute real value added one must have data on the real quantity of intermediate inputs, not just the value of intermediate inputs.

One data set that has the necessary information for a set of countries is EUKLEMS.<sup>28</sup> We begin, therefore by using the EUKLEMS data starting in 1970 to compute TFP in the production of value added in agriculture, manufacturing, and services for the same set of countries as in Section 2: Australia, Canada, Japan, Korea and the USA as well as the aggregate of 15 EU countries. Figure 10 plots the sectoral TFP's for these countries. Given that we are interested in growth rates of TFP, we normalize TFP in the initial year for all sectors in all countries to be one. One message is that there are indeed substantial differences in the growth rates of TFP across sectors. Moreover, we can see that the conditions of Ngai and Pissarides (2007) broadly hold for Australia, Canada, the EU 15, and the United States: averaging over the time period

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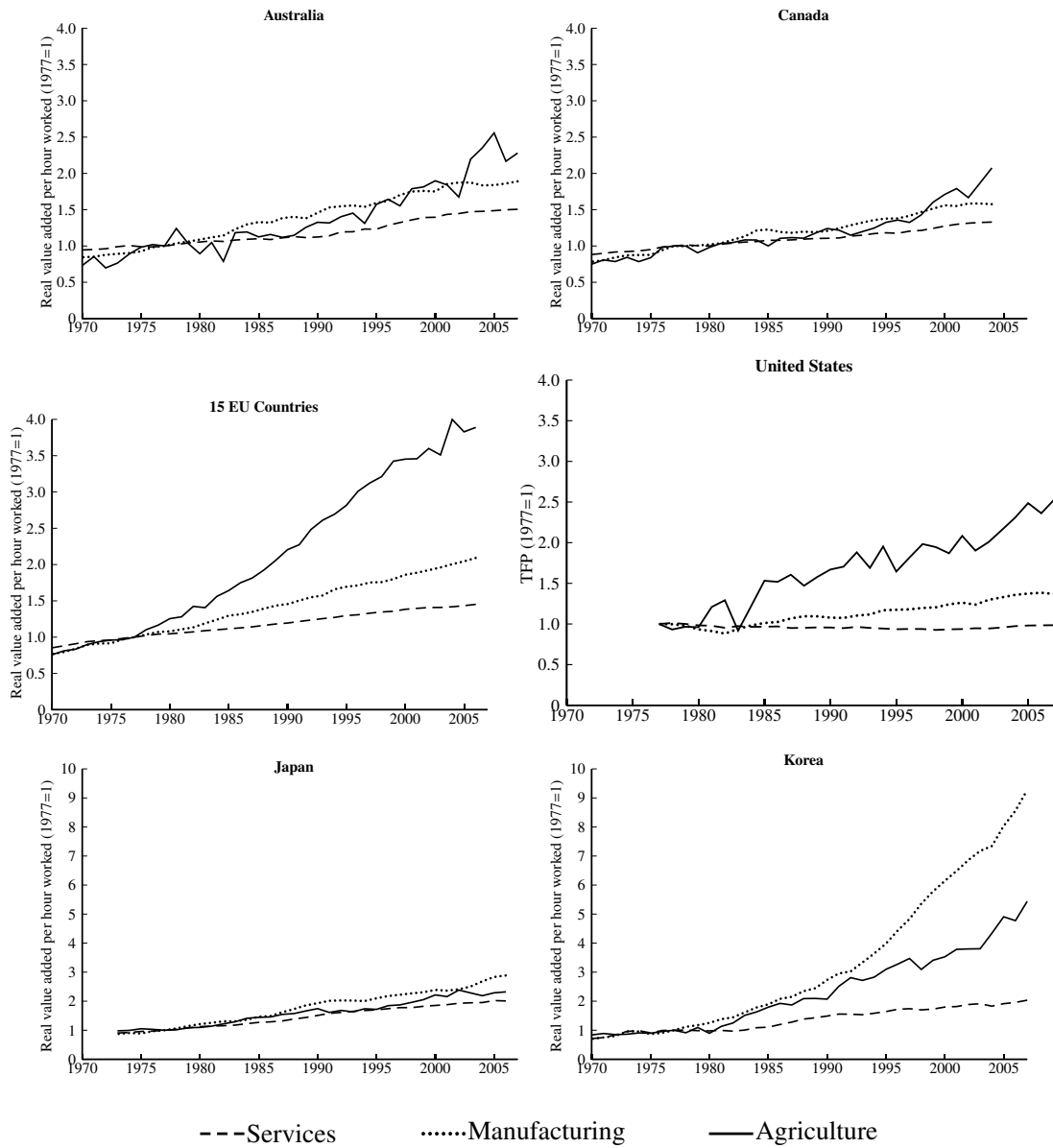
<sup>28</sup>See Timmer et al. (2010), particularly the chapter on structural change, for further discussion of the details of the EUKLEMS data on multifactor productivity. See also Duarte and Restuccia (2010) who document similar facts about TFP as we do here.

1970–2007, TFP in agricultural shows the strongest growth while TFP in services show the weakest growth. This is exactly what is needed for the observed reallocation of employment out of agriculture and manufacturing into the service sector in the model of Ngai and Pissarides (2007).

While data limitations make it difficult to obtain long time series evidence on sectoral TFP for a large sample of countries, our theory suggests an alternative method which requires fewer data. Specifically, in the analysis of our benchmark model we highlighted the fact that if sectoral production functions are Cobb–Douglas with equal capital shares then there is a direct inverse relationship in equilibrium between changes in relative prices and changes in relative productivities. Given appropriate data on prices, one could use this relationship to infer changes in relative productivity. Since long time series of price data is much more readily available than the data needed to measure TFP directly, this is an appealing alternative. However, in addition to requiring the assumption of Cobb–Douglas production functions with equal capital shares, there are two limitations to be noted. First, in our model we assumed that technological change was the only factor that varied over time. One can easily imagine policies or regulations that may also affect relative prices across sectors. If these factors are important for some countries during some periods, it may be misleading to assume that all relative price changes are driven by changes in relative productivities. Second, although price data do exist going quite far back in time, the price data that is required to infer relative productivity growth in *value added* production functions is the price per unit of *value added*. In practice, most available price indices do correspond to final goods or to gross output.

Having noted these qualifications, we turn to the evidence documented by Alvarez–Cuadrado and Poschke (2011) about time series changes in the relative price of agriculture to non–agriculture for eleven advanced countries over the last two centuries. A key feature of these data is that the price of agriculture relative to non–agriculture changed its behavior during the last two centuries: while before World War II it showed an increasing trend, after World War II it started to follow a decreasing trend. Interpreting these changes in relative prices as indicative of changes in relative TFPs, the implication is that prior to World War II, TFP growth in agriculture was actually lower than in non–agriculture. The period before World War II also

**Figure 10: Sectoral TFP for Selected Countries  
– Time Series from EUKLEMS 1970–2007**



**Source:** EUKLEMS

corresponds to the period that saw the largest movement out of agriculture. In contrast to the findings for data since 1970, the longer time series does not seem to be consistent with relative TFPs driving the labor reallocation from agriculture to non-agriculture.

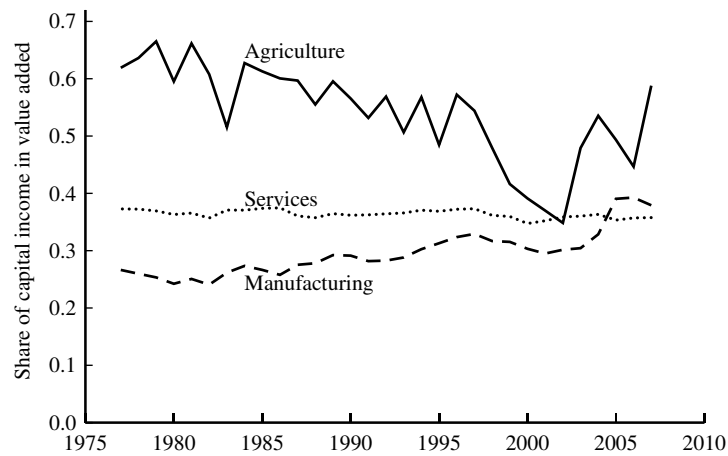
By way of summary, we think there are two main conclusions that can be drawn from this evidence. First, there are systematic differences in TFP growth rates across sectors. After World War II these differences appear to be consistent with what is needed to obtain the observed reallocation of employment out of agriculture and manufacturing into the service sector in the model of Ngai and Pissarides (2007). Second, the differences in TFP growth rates across sectors do not appear to be stable over very long periods of time, at least in the case of agriculture versus non-agriculture, which does not bode too well for either of the models of structural transformation and exact balanced growth that we highlighted previously.

### **5.1.2 Differences in Sectoral Capital Shares**

Next we consider evidence regarding the potential role of differences in sectoral capital shares, as emphasized by Acemoglu and Guerrieri (2008). The specific goal is to assess whether condition (45) is borne out by the data, we calculate the US capital shares in value added at the sector level for the period 1977–2007. We follow the method of Gollin (2002) and first calculate the capital shares in value added without proprietors' income; we then split proprietors' income between capital and labor in the proportion that we found for value added without proprietors' income. In order to avoid confusion, we stress that these capital shares refer to value added at the sector level, and not to final expenditure. The capital shares for final expenditure at the sector level can be found in a related paper, Valentinyi and Herrendorf (2008), and in general they are different from those for value added.

Figure 11 reports the results. We can see that on average the capital share in agriculture is considerably higher than in the other two sectors, but that the capital share in manufacturing is lower than the capital share in services. Hence, differences in capital shares cause reallocation from agriculture to manufacturing and from services to manufacturing. The second effect runs counter to what we have documented above.

**Figure 11: Sectoral Capital Shares – USA 1977–2007**



**Source:** EUKLEMS

## 5.2 The Importance of Substitution and Income Effects

Since the theoretical literature has emphasized income and substitution effects, it is natural to ask what the data say about these two effects. There are two natural and complementary approaches to this question. In the spirit of our earlier analysis, one approach starts with a stand-in household and uses aggregate data to infer the relative importance of the two different mechanisms. The second approach uses data on individual households to estimate properties of preferences and then assesses the implications for aggregate behavior. In the interest of space, we will focus on the first approach, though we will briefly mention some results from the analysis of micro data. We discuss two recent contributions: Dennis and Iscan (2009) and Herrendorf et al. (2009). The former studies the forces leading to the movement of activity out of agriculture in the United States over the last two centuries, whereas the latter focuses specifically on the reallocation of activity across all three sectors in the United States since 1947. We describe each in turn.

### 5.2.1 The Movement Out of Agriculture in the US Since 1800

Dennis and Iscan (2009) seek to assess the relative importance of income effects, relative TFP growth and capital deepening on the movement of labor out of agriculture in the US over the

last two centuries. Their framework is very similar to our benchmark model with the exception of three details. First, they have only two sectors, agriculture and non-agriculture. Second, they assume that all investment comes from the non-agricultural sector. Third, they do not impose that the capital share is the same in both sectors. Initially, Dennis and Iscan write the utility function as the two-sector analogue of our utility function, but in their empirical analysis they also allow for the possibility that the subsistence term  $\bar{c}_a$  changes over time. Given our earlier discussion, we note that while this general specification is not consistent with generalized balanced growth, but it captures the basic forces that the theoretical literature has emphasized.

Dennis and Iscan (2009) derive an equilibrium relationship that expresses the share of labor devoted to agriculture as a function of three factors, which in turn reflect income effects through the subsistence term, relative productivity effects via differential growth rates of TFP, and capital deepening effects. Expressed in terms of our notation, this equilibrium relationship is:<sup>29</sup>

$$1 - n_{at} = \frac{1 - s_a(c_{at})}{1 + p_R(A_{at}, A_{nt})s_k(k_{at}, k_{nt})s_X(c_{nt}, X_t)} \quad (47)$$

where

$$\begin{aligned} s_a(c_{at}) &= \frac{\bar{c}_a}{c_{at}}, & p_R(A_{at}, A_{nt}) &= \frac{\omega_a}{\omega_n} \left( \frac{A_{nt}}{A_{at}} \right)^{1-\varepsilon}, \\ s_k(k_{at}, k_{nt}) &= \left( \frac{1 - \theta_a}{1 - \theta_n} \right)^\varepsilon \left( \frac{k_{nt}^{\theta_n}}{k_{at}^{\theta_a}} \right)^{1-\varepsilon}, & s_X(c_{nt}, X_t) &= \frac{X_t}{c_{nt} + X_t}. \end{aligned}$$

The term  $1 - s_a(c_{at})$  captures the income effect that operates through the subsistence term  $\bar{c}_a$ . The terms  $p_R(A_{at}, A_{nt})$  and  $s_k(k_{at}, k_{nt})$  capture the relative price effects that arises from differential technological progress and capital deepening, respectively, while the term  $s_X(c_{nt}, X_t)$  captures the effects associated with changes in the investment rate.

Dennis and Iscan (2009) calibrate the key parameters of the model (elasticity of substitution, subsistence terms, preference weights, and capital shares) and then assess the extent to which equation (47) holds in the data. In particular, they substitute actual values into the right-hand side of equation (47), solve for the implied share of labor allocated to agriculture and compare this to the actual series from the data. To assess the importance of the different factors they

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<sup>29</sup>We use the index  $n$  for the non-agricultural sector.

carry out the same exercise but only allowing one of the factors to change over time.

Dennis and Iscan (2009) main findings are as follows. First, the model does a reasonable job of capturing the time series changes in the employment share of agriculture since 1800. If the value of  $\bar{c}_a$  is held fixed throughout, the model somewhat under-predicts the employment share for agriculture in the 1800s, but does fine in the post 1950 period, although a small time trend over the period 1800–1950 yields a better fit over the entire period. Second, prior to 1950 the income effect is the dominant factor in accounting for the movement of employment out of agriculture, whereas the relative productivity effect is working in the opposite direction. Only in the post 1950 period do the relative productivity and capital deepening effects play even a modest role in accounting for the change in the employment share of agriculture. They also consider various extensions to their analysis, such as incorporating trade and show that the results are robust to these extensions.

We want to stress three key implications of the results of Dennis and Iscan (2009). First, the fact that their model does a reasonable job of capturing the movement of labor out of agriculture over a long time period suggests that our benchmark model is sufficiently rich to capture some key features in the data. Second, the fact that a time varying subsistence term,  $\bar{c}_{at}$ , improves the model's ability to account for the movement out of agriculture is notable, and suggests that a deeper theory of how income effects arise may be warranted. Third, at least for the movement of labor out of agriculture in the United States, income effects are effectively the sole driving force behind this decline; even though the other factors play a role after 1950, this occurs when almost all of the decline in the employment share for agriculture has already happened.

It is also relevant to note some limitations of the analysis in Dennis and Iscan (2009). First, it only focuses on the movement of labor out of agriculture and does not address the issue of what forces shape the allocation of employment between manufacturing and services. Second, all of their results come from a calibration exercise, but there is little direct evidence on some of the key parameters they use for this exercise. Additionally, they connect their model to the data in a somewhat inconsistent fashion, in that they interpret their production functions as value added production functions, but when they look at consumption of agriculture they interpret it as consumption of final goods. In the next subsection, we discuss in detail why this

is inconsistent. Third, they focus only on the changes in employment shares, and so do not address the issue of the discrepancy between value added shares and employment shares that we documented earlier. Nonetheless, we think that this paper makes an important contribution to the effort to identify the key economic forces behind structural transformation.

A related exercise was carried out by Buera and Kaboski (2009). Specifically, they assessed the ability of a calibrated version of our benchmark (three-sector) model to account for the broad patterns of structural transformation in the US from the 1800s to the present. One difficulty that they noted was the ability of the model to account for the acceleration in the nominal value added share of the service sector in the post World War II period.

### 5.2.2 Structural Transformation in the US Since 1947

Herrendorf et al. (2009) offer a related but distinct approach to uncovering the importance of income and substitution effects in accounting for structural transformation. In contrast to Dennis and Iscan (2009), who considered the allocation of employment between agriculture and non-agriculture in the US since 1800, Herrendorf et al. (2009) consider the reallocation among consumption expenditure shares for all three sectors in the US since 1947. Specifically, starting with a stand-in household, they asked whether the utility function in (1) provides a good fit to the US data on expenditure shares in the post World War II period, and if so, what this implies for the values of the key parameters  $\bar{c}_a$ ,  $\bar{c}_s$  and  $\varepsilon$ , and the implied importance of income and substitution effects.

Although this seems to be a simple question, Herrendorf et al. (2009) argued that the question is not even properly specified. The reason for this is related to the difference between value added and final expenditure, which we have previously discussed. In particular, if one interprets the sectoral production functions as value added production functions then the arguments of the utility function necessarily represent the corresponding consumption of sectoral value added. In terms of our previous example of the purchase of a cotton shirt, this implies that the shirt is broken into three value added pieces, each of which the household values as they contribute to the three different categories of value added. Herrendorf et al call this the *value added approach*. Alternatively, one may interpret the commodities in the utility function as



final expenditure categories, as is typically done in household expenditure studies. The outputs of the production functions must then be viewed as final expenditure rather than value added. In terms of the purchase of a cotton shirt, the consumer simply derives utility from the shirt as a whole as it contributes to the single category of manufacturing consumption. Herrendorf et al call this the *final expenditure approach*. It is important to note that there is no right or wrong in terms of these two approaches. From the perspective of preferences, these are simply two different ways of aggregating across the many characteristics that consumers value. As is true with any attempt to aggregate individual characteristics into broader groups, one can imagine examples where one approach seems preferable.

The choice of interpretation matters if the relative prices and quantities are not the same for the two different interpretations. In particular, even if the two different approaches display similar qualitative properties in terms of changes over time, differences in quantitative properties may have important implications for parameters of the utility function and the importance of income and substitution effects. Herrendorf et al. (2009) carry out the manipulations necessary to have consistent sets of data for the two approaches and they provide the following answers.

One possible outcome from this exercise is that one of the approaches provides a better fit to the data, in which case one might use this as evidence in support of one approach over the other. However, Herrendorf et al. (2009) found that for both approaches the preferences represented by (1) yield very good fits to the postwar US data on relative prices and expenditure shares. However, the two approaches yield very different parameter estimates for the utility functions and very different assessments of the relative importance of substitution versus income effects.

For the final expenditure approach, income effects are the dominant source of changes in expenditure shares, and the Stone–Geary utility function (28) of Kongsamut et al. (2001) provides a good fit to the data.<sup>30</sup> For the value added approach, it turns out that substitution effects are the dominant source of changes in expenditure shares. In particular, the homothetic Leontief utility function  $\min_{c_{at}, c_{mt}, c_{st}} \{\omega_a c_{at}, \omega_m c_{mt}, \omega_s c_{st}\}$ , which results for  $\varepsilon = \bar{c}_a = \bar{c}_s = 0$ , provides a good fit to the data. Interestingly, this utility function is a special case of the class of inelastic

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<sup>30</sup>Many other papers have estimated linear expenditure systems implied by the Stone–Geary utility specification. A review of this literature is Blundell (1988).

CES utility functions that Ngai and Pissarides (2007) considered.<sup>31</sup>

It is important to emphasize what these results mean. In particular, these results are not an example of researchers obtaining different estimates for a given parameter from different data sets, suggesting that further work is needed to narrow down the set of possible values. Instead, the implication is that there are two different ways to interpret commodities in the utility function in multi-sector models. It turns out that being explicit about which interpretation is adopted is of critical importance, in that it has implications for what data is required to connect the model with the data, and as just shown, this has very important implications for implied preference parameters. Furthermore, note that the two approaches are just two different aggregate representations of the same underlying economic data. The key message is that what one cannot talk about the importance of income or substitution effects as drivers of structural transformation without specifying what representation of the data one is adopting. What shows up as income effects in one representation may manifest itself as substitution effects in a different representation. Different representations are connected via the complex input-output relationships in the economy. Herrendorf et al. (2009) show how one can construct the mapping between the two representations for a given input-output structure.

We stress two key results. First, the fact that the model is able to account for changes in expenditure shares for the US since 1947 is again support for the parsimonious model that we have adopted as our benchmark. Second, it highlights that empirical researchers working with multi-sector models must take care to be explicit about how commodities in utility functions are to be interpreted. Different interpretations have dramatically different implications for how the models are to be connected with the data and what the implied parameters of the utility function.

One of the limitations of this study is that it only focuses on the post 1947 period for the US, and this is a period in which the US has already experienced much of the reallocation out of agriculture. While it is of interest to extend this type of analysis to longer time periods and different countries, a key issue is data availability.<sup>32</sup>

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<sup>31</sup>In independent work, Buera and Kaboski (2009) also reached the conclusion that a low  $\sigma$  is required to match value added data.

<sup>32</sup>This is relevant for the analysis of Buera and Kaboski (2009). They carry out a calibration exercise for the US over a longer time period, but need to use different sources for relative prices in the pre 1947 period. Given that

## 6 Extensions of the Benchmark Model

In this section we discuss relaxing three features present in the analysis of the benchmark model. The first is the assumption of a closed economy. The second is the assumption of that there is no cost of moving labor across sectors (“perfect labor mobility”). The third is the assumption that there are no costs of moving goods across sectors (“zero transportation costs”).

### 6.1 International Trade

Thus far our theoretical analysis has taken place under the assumption of a closed economy. A key implication of being a closed economy is that the production of each of the four sectors must equal the corresponding household choices (either of investment or of one of the three consumption goods). The equality between sectoral productions and consumption/investment played a key role in generating the results concerning structural transformation that we obtained in the benchmark model. For example, in the model of Ngai and Pissarides (2007), we saw that labor moved out of the consumption sector that had the highest productivity growth because of the household’s desire to maintain the composition of its consumption allocation (inelastic demand). In the model of Kongsamut et al. (2001), technological progress was uniform across sectors, but labor moved out of agriculture because of the household’s desire to change the composition of its consumption allocation towards manufactured goods and services (differences in income elasticities).

In this subsection we discuss the extent to which openness changes the results about structural transformation. We begin with the simple observation that the competitive equilibrium of a model in which all commodities are tradeable without costs will have a complete separation between the decisions of firms and households. This observation implies that in an open–economy version of our benchmark model without trade costs the production measures of structural transformation (i.e. employment and value added shares) would generically follow a different pattern than the consumption expenditure share. This is relevant because, as we have

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prices for value added consumption and final consumption are quite different in the post 1947 period and have very different implications for preference parameters, an issue arises with how to interpret results that use a mixture of prices.

documented in Section 2, there is a discrepancy between production and consumption shares in some instances, most notably for the share of manufacturing in Korea.

Matsuyama (2009) was the first to analytically work out the idea of the previous paragraph for a simple two-country model. He abstracts from capital and considers a Stone–Geary utility function over the three consumption goods food, manufactured goods, and services. He assumes that agricultural goods are an endowment whereas manufactured goods and services are produced with technologies that are linear in labor, and that agricultural and manufactured goods can be traded with the rest of the world at zero trade costs whereas services cannot be traded. Matsuyama shows two results for this simple model. First, if there is technological progress in manufacturing then the total manufacturing labor of both countries declines. Second, if only one country experiences stronger technological progress in manufacturing than the other country, then manufacturing labor in the first country may initially increase while manufacturing labor in the second country decreases unambiguously. Eventually, when technological progress in the manufacturing sector has been sufficiently strong, the share of manufacturing labor in the first country will decrease also. These results suggest that an inverted U shape relationship may occur in the country which experiences the stronger technological progress in manufacturing.

Yi and Zhang (2010) generalize the idea of Matsuyama to a two-country–version of our benchmark model of structural transformation, in which all goods are produced with labor only. The assumption that agricultural and manufactured goods are tradeable without costs would then lead to the counterfactual implication that each country specializes in either agriculture or manufacturing. They therefore assume that each of the three sectors is the aggregate of a continuum of goods as in Eaton and Kortum (2002). Yi and Zhang (2010) simulate their model under the assumption that one country has higher productivity growth in manufacturing than the other country. They provides examples for which the country with the higher productivity growth in manufacturing experiences an inverted U shape in the shares of manufacturing employment and value added while the other country experiences a downward sloping shape in the shares of manufacturing labor and value added.

From the empirical perspective it is of interest to ask whether besides the inverted U shape

of manufacturing employment and value added there is evidence for the effects of openness on structural transformation. One clear prediction of the models of Matsuyama (2009) and Yi and Zhang (2010) is that the labor shares of sectors that produce tradeable goods should differ across countries that have different sectoral productivities. In Section 2 we noted that there was some evidence of dispersion in sectoral labor shares across countries in the European Union and Japan, with Germany and Japan having unusually large share of manufacturing hours worked and Korea having an unusually large share of real manufactured value added. Betts et al. (2011) study the role of international trade in Korea's industrialization in a two-country model with three sectors. They find that international trade played a crucial role for the rapid rise in the manufacturing value added and employment shares, but that it did not play much of a role for the decline of Korean agriculture. While such story may be consistent with various accounts regarding the importance of trade in the development of Korea, it is hard to reconcile with the patterns we found in Section 2. Specifically, we found there that Korea did not display any distinctive behavior for the nominal value added share in manufacturing.

From a theoretical perspective, we conclude that to the extent that openness matters for studying structural transformation, it is most likely to show up in a discrepancy between value added and consumption in sectors that trade with the rest of the world. In the past, this applied to manufacturing, and to a lesser extent to agriculture. In recent years, however, there has been an increasing trend toward trade in services. An open question moving forward concerns the extent to which increased trade in services will exert an influence over the nature of structural transformation. For example, will this hasten the movement of resources out of manufacturing in a country like the US which is thought to have relatively high productivity in many service industries?

## **6.2 Labor Mobility**

Our benchmark model assumed that labor was homogeneous and could be allocated across sectors without any labor mobility costs. There are several interesting issues that arise when there labor mobility costs. In this subsection we discuss the most relevant ones.

We begin with the paper by Lee and Wolpin (2006) about the large reallocation of labor

from manufacturing to services in the United States over the period from 1968 to 2000. The goals of this paper are to measure the costs associated with sectoral labor reallocation and to assess the relative importance of labor demand and supply factors for sectoral labor reallocation, where labor demand factors are defined as changes in sectoral productivity and relative prices and labor supply factors are defined as changes in demographics, fertility, and educational attainment. To reach these goals, they develop a framework with a detailed labor market. To begin with, there are three occupational choices in each sector: blue collar, white collar, and pink collar (i.e., secretarial, clerical etc...). Moreover, workers differ in their educational attainment and they can accumulate sector-specific and occupation-specific human capital while working. Lastly, there are various types of technological change and the production functions have a constant elasticity of substitution between capital and labor.

Lee and Wolpin (2006) estimate their model using micro data. Their main findings are as follows. First, labor demand factors are the key driving forces behind the reallocation of labor across sectors. In contrast, labor supply factors do not play much of a role. This finding is consistent with the emphasis that our benchmark model puts on technological factors. Second, and in contrast to our benchmark model, the mobility costs associated with moving across sectors are large; for example, the monetary cost of changing sectors can be as large as 75 percent of annual earnings. Moreover, changing occupations within a sector is significantly less costly than changing sectors while maintaining the same occupation.

Lee and Wolpin (2006) carry out several counterfactuals regarding how changes in mobility costs would have affected the evolution of labor market outcomes. Interestingly, they find that if mobility costs had been zero, aggregate productivity would have been higher and the labor market histories of individual workers would have been different, but the evolution of sectoral employment shares and value added shares would not have changed much. The economics behind this result is that with lower mobility costs workers can better allocate their time to the sector in which their idiosyncratic productivity is highest. This raises aggregate productivity and changes the labor market histories of individual workers. However, since it leads to flows of workers in *both* directions, the effect on relative sectoral employment is relatively small. This result suggests that abstracting from mobility costs in our benchmark model does not have

large quantitative effects on the sectoral employment allocation.

Lee and Wolpin (2006) also ask what would have happened if sectoral labor mobility had been more costly. They find that while there would have been little effect on trend changes in employment shares, the level of the employment share of services would have shifted upward. This result runs counter to the intuition that increased mobility costs will decrease the flow of workers into the expanding service sector. To understand this, it is important to realize that this intuition is based on how mobility costs affect the response to an unanticipated shock. In contrast, what matters for Lee and Wolpin's exercise are the choices that forward looking new entrants make in the face of the trend that the service sector is becoming more attractive in comparison to the goods sector. If we increase the size of mobility costs, then more entrants move directly into the service sector, instead of first going to the manufacturing sector and later switching to the service sector.

There is more evidence that the role of new entrants is crucial for the labor reallocation across sectors in the context of structural transformation. For example, Kim and Topel (1995) show that during Korea's rapid industrialization almost all of the changes in the sectoral employment shares of agriculture and manufacturing resulted from changes in the behavior of new entrants. As a result, the large decrease in the agricultural employment share and the large increase in the manufacturing employment share were accomplished with little reallocation of existing workers.<sup>33</sup> To the extent that new entrants are an important source of labor market flexibility one might conjecture that economies with different rates of growth in the labor force might experience different patterns of structural transformation. However, we are not aware of existing evidence that supports this conjecture.

While some mobility costs might reflect technological factors, there is also the possibility that policies, regulations and institutional factors lead to the barriers to labor mobility. Examples include implicit or explicit firing costs levied on employers, subsidies to establishments in declining industries, entry barriers that make it costly for firms to start up new establishments, generous unemployment benefits or early retirement schemes that are offered to displaced workers, and direct restrictions on the mobility of workers.<sup>34</sup> There are many studies of these types

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<sup>33</sup>Matsuyama (1992) and Rogerson (2006) both present models of sectoral reallocation that have this property.

<sup>34</sup>China is a clear example of an economy that has direct restrictions on the mobility of workers, though we are

of factors, but most of them make no reference to the process of structural transformation. The reason for this is that most job creation and destruction occurs within rather than across narrow industrial classifications, and so the main effects come from the reallocation of resources across establishments when jobs are created and destroyed.

Three exceptions that study the effects of labor mobility costs in the context of structural transformation are Nickell et al. (2002), Messina (2006) and Hayashi and Prescott (2008). Nickell et al. (2002) examine the correlations between the sectoral composition and various policy and institutional factors in a panel data set panel of 14 OECD countries and 5 one-digit industries during the period 1975–94. One of their findings is that countries with more stringent employment protection policies have larger industrial sectors, suggesting that employment protection policies might impede the reallocation of employment from manufacturing into services. Messina (2006) considers the role of entry barriers. One distinguishing feature of structural transformation in Europe is that condition on aggregate productivity (i.e., output/hour), Europe has a much lower employment share for services than do other rich countries.<sup>35</sup> Messina argues that this is the result of higher entry barriers in Europe, including such factors as direct costs associated with licensing and indirect costs associated with zoning restrictions or regulations that restrict shopping hours, etc. Because the reallocation of workers into services requires additional entry of establishments into the service sector, these barriers retard the movement of economic activity into the service sector. Hayashi and Prescott (2008) study the movement of labor out of agriculture in Japan before World War II. They argue that the prewar patriarchy that forced the son designated as heir to stay in agriculture effectively amounted to a barrier to the movement of labor out of the agriculture sector. Using a standard neoclassical two-sector growth model, they show that the barrier-induced sectoral distortion, and the implied lack of capital accumulation account well for the depressed output level of Japan's prewar economy.

Although Lee and Wolpin (2006) incorporated a range of factors that make mobility costly for individual workers, their model still shared the feature of our benchmark model that all labor reallocation was voluntary from the perspective of the worker. A large literature has

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not aware of any studies that have assessed the impact of this.

<sup>35</sup>This was not apparent in Section 2 since we plotted the service share of hours worked versus per capita income rather than output per hour.



documented the large earnings losses that older workers face when they are displaced; see, for example, Jacobson et al. (1993). To many policymakers and commentators, the reallocation of labor from manufacturing to services that is part of the process of structural transformation is synonymous with the displacement of older, high-tenure workers in the manufacturing sector and either unemployment or large losses in earnings. While the connection may seem clear cut, direct evidence on this point is much less clear cut. As noted earlier, most job creation and destruction occurs within narrow industry classifications, and so is not directly related to the reallocation of activity across broad sectors.

### **6.3 Goods Mobility**

If openness matters for the process of structural transformation in some settings then it follows that the cost of moving goods may will presumably influence structural transformation as well through their effect on trade. More interesting is the possibility that transport costs might influence structural transformation in a closed economy setting. One simple idea in this literature stems from noting that while agriculture is predominantly rural, much of the activity outside of agriculture takes place in cities. It follows that food consumed by non-agricultural workers needs to be transported from rural to urban areas. If this is the case then high costs of moving food from rural areas could exert a negative influence on the movement of labor out of agriculture.

Gollin and Rogerson (2010) formalize this in the context of simple two-sector static model that focuses on the allocation of labor between agriculture and non-agriculture. Consistent with our benchmark model, their model also allows for both income effects via a subsistence term in the utility from agriculture, and productivity effects in terms of the factors that determine the allocation of labor to agriculture. They carry out some numerical exercises to suggest that transportation costs can exert an important influence on the allocation of labor across sectors. An interesting feature of their numerical examples is that there is a strong interaction between increases in income and improvements in transportation costs in terms of their impact on labor moving out of agriculture.

## 7 Applications of Structural Transformation

In this section we return to the question we posed in the introduction to this chapter: Does incorporating structural transformation into the standard growth model deliver new insights? In other words, is there a substantive payoff to working with versions of the growth model that account for structural transformation? We discuss several issues where changes in the sectoral composition of the economy matter has been shown to matter. We conclude that explicit modeling of structural transformation then offers important additional insights.

### 7.1 Structural Transformation and Economic Development

Caselli (2005) and Restuccia et al. (2006) argue that the proximate cause of much of the large differences in living standards across countries is attributable to two simple facts: (1) developing countries are much less productive in agriculture relative to developed countries, and (2) developing countries devote much more of their labor to agriculture than do developed countries. These two facts suggest that in order to understand why developing countries are so poor it is of first-order importance to understand the forces that shape the allocation of resources between agriculture and the other sectors. A version of the growth model extended to incorporate structural transformation is the natural framework to be used in this context.

Work by Gollin, Parente, and Rogerson (2002, 2007) illustrates how low agricultural productivity can be the source of large cross-country differences in aggregate productivity. For ease of exposition we focus on the simpler presentation in the 2002 paper, which uses a two-sector version of our benchmark model, with the two sectors being agriculture and non-agriculture. They assume that the population is constant and normalize it to one. Preferences are such that there is a subsistence level  $\bar{c}_a$  of agricultural consumption at which individuals are also satiated. The non-agricultural production function is essentially a Cobb–Douglas production function in capital and labor. In contrast, there are two agricultural production functions: a traditional and a modern one.<sup>36</sup> Both agricultural production functions are linear in labor, though the analysis would be unaffected by assuming a fixed quantity of land and decreasing returns to scale in la-

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<sup>36</sup>Hansen and Prescott (2002) use a similar assumption but at the aggregate level.

bor. The traditional production is assumed to be the same across countries and to be sufficiently productive to exactly meet subsistence agricultural needs when all labor is allocated to it. The modern production function has a country-specific TFP parameter and it is the only production function that is subject to technological progress.

In this model, only the agricultural technology with the larger productivity will be used in equilibrium. Initially this is the traditional technology. Since the modern technology is subject to technological progress, at some point the modern technology will replace the traditional technology as the only technology that will be used. The somewhat extreme structure of the model then yields a very simple solution method for determining the equilibrium. Total food production must be  $\bar{c}_a$ . As long as the traditional technology is used, this means that all labor will be in agriculture. When the modern technology starts to dominate the traditional technology, labor will start to flow from agriculture to non-agriculture. With the time series for labor allocations determined, the remainder of the model becomes a standard growth model with an exogenously given process for labor. The growth rate of labor in the non-agricultural sector is completely determined by the exogenous growth rate of labor productivity in the modern agricultural sector. Since all countries have the same output of agriculture, cross-country differences in aggregate output are entirely driven by differences in non-agricultural output.

Several implications follow. First, countries that use the modern technology in agriculture but have low productivity in it will have to devote more labor to agriculture. This leads to less labor, and capital, in non-agriculture, and hence to less aggregate output. Given the observed differences in the amount of labor that is devoted to agriculture, show that this mechanism can account for a large part of the cross-country differences in aggregate output. This is interesting because in their model the only difference across countries is the level of productivity of agriculture.

Second, assuming that productivity growth rates are constant over time, the model necessarily implies that transition dynamics will be long-lived, thereby addressing a point emphasized by King and Rébello (1993) that in a standard one-sector growth model transition to the steady state capital level is rapid.<sup>37</sup> This point does not carry over to the two-sector model because la-

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<sup>37</sup>Chang and Hornstein (2011) make a related point about Korea. They show that two modifications of the one-sector growth model are essential to account for the long-lived transition dynamics since 1960 during which

bor allocated to the non-agricultural sector only slowly converges to its asymptotic level. Third, the model implies that (in a closed economy setting) advances in agricultural productivity are a precondition for growth. This view was a central argument of Schultz (1953), and figured prominently in later contributions by Johnston and Mellor (1961), Johnston and Kilby (1975), Timmer (1988), and Yang and Zhu (2009), among others. More recently, it has taken a central state in the writing of non-economists such as Diamond (1997).<sup>38</sup>

Laitner (2000) considers a similar framework as Gollin et al. (2002) but focuses on a different issue. He notes that in the time series data there is evidence of an increase in savings rates early in the industrialization process. Whereas some have argued that the increase in savings rate is the driving force behind the industrialization process, Laitner shows that, in a model of structural transformation, this apparent increase in savings rate is simply an artifact of how NIPA measures saving. Early in the development process most labor is employed in agriculture, and so most savings take the form of realized capital gains in the value of land, which is not recorded as savings by the NIPA. As labor moves out of agriculture and agriculture becomes a smaller part of aggregate output, this issue becomes less important quantitatively. Laitner argues that viewed from the perspective of his model of structural transformation, one should not attach any significance to the apparent increase in savings rates that occur in the early stages of development.

## **7.2 Structural Transformation and Regional Income Convergence**

One of the dramatic secular changes in the US economy over the post World War II period is the convergence of incomes across regions; see, for example, Barro and Sala-i-Martin (1992). In the context of standard one-sector neoclassical growth model, this convergence in incomes would be attributed to changes in either regional TFP or regional factor accumulation. Caselli and Coleman (2001) show that a model of structural transformation provides a richer understanding of the economic forces at work. The motivation for their analysis is provided by the

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Korea continued to accumulate capital. The first essential modification is to distinguish between agriculture and non-agriculture and to take into account that Korean agriculture used relatively little physical capital. The second essential modification is to model that the relative price of capital remained high during most of the transition dynamics.

<sup>38</sup>See Tiffin and Irz (2006) for a recent empirical assessment.

fact that the convergence in regional incomes between the North and the South of the United States coincided with a dramatic narrowing of regional differences in the employment share in agriculture. To carry out their analysis they modify our benchmark model along several dimensions. First, they consider a two–sector version of the model, with the two sectors being agriculture and non–agriculture. Second, they consider a two–region version of the model, where each region has the same structure as our model and there is free mobility of goods across regions. Production opportunities in non–agriculture are the same across regions, but the South has a higher TFP in agriculture. Land is included as a factor of production in agriculture. Third, they assume that there are mobility costs in terms of sectoral reallocation of labor. Specifically, all workers begin in the agricultural sector, and they must pay a cost if they are to move to the non–agricultural sector. They interpret this mobility cost as the cost of acquiring skills that are needed in the non–agricultural sector and argue that it is necessary if one is to account for the secular changes in labor allocations and relative wages.

The basic economics of their analysis is the following. When the United States was relatively poor, more of its workers were engaged in agriculture, due to non–homothetic preferences which imply a large share for agricultural expenditures at low levels of income. Because the South had a comparative advantage in agriculture, the South was doing relatively more agriculture. Because of mobility costs, wages were higher in non–agriculture. The fact that the South was more heavily involved in agriculture therefore led to lower incomes in the South. Over time, production technology in non–agriculture advanced, leading to a decline in the share of workers in agriculture. They also posit that mobility costs decreased, therefore leading to convergence between agricultural and non–agricultural wages.

### **7.3 Structural Transformation and Aggregate Productivity Trends**

Our model of structural transformation allows for the possibility that different sectors have different levels as well as growth rates of labor productivity. Herrendorf and Valentinyi (2011) provide evidence from the 1996 Benchmark Study of the Penn World Tables on sectoral TFP differences across countries. They find that there are large sectoral TFP differences relative to the United States not only in agriculture, but also in manufacturing, and that the sectoral TFP

differences in these two sectors are much larger than in the service sector. Aggregate labor productivity may then be affected by the sectoral composition of the economy. In particular, to the extent that different countries are at different stages of the process of structural transformation, sectoral reallocation associated with structural transformation could generate significant changes in aggregate productivity growth [Echevarria (1997)]. In principle, episodes of acceleration or slowdown in aggregate productivity growth may occur even if in each country sectoral productivities are growing at constant rates.

In a recent paper, Duarte and Restuccia (2010) have investigated the importance of these effects in a sample of 29 countries for the period of 1956–2004. They employed a somewhat simplified version of our benchmark model in which labor is the only factor of production (and production functions are linear in labor). They assumed that each sector’s labor productivity grows at a constant rate, but that level and growth rates differ across economies as dictated by the data.

The preference structure of Duarte and Restuccia (2010) assumes a period utility function which is a two–period version of (28):

$$C_t = \omega \log(c_{at} - \bar{c}_a) + \omega_n \log(c_{nt})$$

$c_{nt}$  stands for non–agricultural consumption and it is a CES aggregator of manufactured goods and services. Preference parameters are calibrated so as to match the behavior of the US economy and are assumed to be the same across countries. The initial productivity levels of all countries relative to the US are inferred from the model by requiring that the model match the observed employment shares in the initial period. Inputting the sectoral productivity growth rates from the data, Duarte and Restuccia (2010) then simulate the model and compute the implied series for aggregate labor productivity.

Even though their model assumes constant productivity growth rates at the sectoral level of each country, it generates large movements in relative aggregate productivity across countries over time. Key to this finding is that differences in the levels and growth rates of labor productivity between rich and poor countries are larger in agriculture and services than in man-

ufacturing. This implies that during the process of structural transformation, the reallocation of labor from agriculture to manufacturing leads to a catch up of aggregate productivity relative to the USA, and the reallocation from manufacturing to services leads to a falling behind of aggregate productivity relative to the USA.

In related research, Bah and Brada (2009) study the countries from Central Europe which have recently entered the European Union. The point of departure of their analysis is the stylized fact that central planning during communist times resulted in “over-agrarianism” and “over-industrialization”, and the neglect of service sector in these countries. Bah and Brada document that even today employment in the service sector is considerably smaller in Central Europe than in the core countries of the European Union. Moreover, they find that in all of these countries the service sector has lower TFP than the manufacturing sector. This implies that structural transformation into the service sector will lead to losses in GDP per capita, unless reforms are implemented that make the service sectors more productive.

#### **7.4 Structural Transformation and Hours Worked**

Following Prescott (2004), there is a sizeable literature that seeks to understand the large differences in hours worked that have emerged over time between the USA and countries in continental Europe. Prescott used the standard one-sector growth model to demonstrate that changes in labor taxes could account for much of the emerging difference.

Rogerson (2008) argued that a model of structural transformation provides additional insights into the evolution of hours dynamics. In particular, he compared the evolution of hours worked in the United States to those in an aggregate of five continental European economies (Belgium, France, Germany, Italy, Netherlands) since 1956. Whereas aggregate hours worked were about 5% higher in Europe in 1956, by 2003 they were more than 30% lower. Looking at the sectoral evolution of hours worked reveals an interesting pattern. During the period in which hours worked in these European economies fell by more than 35% relative to the US, one observes that the relative level of hours worked in the goods sector in Europe fell dramatically, whereas the relative level of hours worked in services remained relatively flat. One might be tempted to conclude that the key to understanding the relative decline in hours worked in Europe

lies in understanding the relative decline in hours worked in the goods sector. However, when one views the sectoral evolution of hours worked in the context of structural transformation one is lead to exactly the opposite conclusion. Specifically, in 1956 Europe was considerably behind the USA in terms of development, and consistent with our earlier empirical analysis, had a much larger share of hours in the goods sector and a smaller share in the service sector than the United States. By 2000 Europe has basically caught up to the United States in terms of productivity. Holding all else constant, one would expect that the sectoral hours worked distribution in Europe in 2000 would look similar to that in the United States. That is, the process of structural transformation leads us to expect that while hours in the goods sector in Europe should have decreased relative to the US, hours in the service sector in Europe should in fact have increased. Put somewhat differently, the issue of understanding why hours worked are so much lower in Europe reduces to the issue of understanding why the European service sector has failed to grow like its counterpart in the US. In fact, this dynamic was apparent in the hours plots in Figure 2.

In addition to simplifying the analysis by aggregating agriculture and manufacturing to one category and by abstracting from capital, Rogerson modified our benchmark model along two key dimensions: he added a labor supply decision and he allowed for home production, which he assumed to be substitutable with the output of the service sector. His model combines both income and price effects to generate structural transformation. Taking changes in productivity and labor taxes as given, he calibrated the preference parameters so as to match the changes in the US economy between 1956 and 2003, including the change in time devoted to home production.<sup>39</sup> He then fed in European values for productivity and taxes in both 1956 and 2003 and examined the ability of the model to account for aggregate and sectoral observations in Europe in 1956 and 2003. Overall, Rogerson found that the model accounts well for the sectoral European labor allocations.

Rogerson assumes that the utility function is non-homothetic in that it has a subsistence level of goods consumption. This turns out to be important for understanding relative hours worked in Europe in the initial year of his study, 1956. At that time, Europe already had higher

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<sup>39</sup>See Aguiar and Hurst (2007) and Ramey and Francis (2009) for evidence on the decline of home production time in the USA.



tax rates than the US, yet they had higher hours of work. The non-homotheticity acts like a negative income effect, and this effect is larger the lower is aggregate productivity. Given that Europe lagged the US in aggregate productivity in 1956, this effect serves to increase hours in Europe relative to the US. Additionally, because the model generates a structural transformation, Europe devoted more labor to goods production than the US in 1956. Because there are fewer non-market substitutes for goods, this effect also serves to increase the amount of time devoted to market work.

In related work, Ngai and Pissarides (2008) add a home production sector to their earlier model of structural transformation that we have discussed above, Ngai and Pissarides (2007). They showed that over time the model with home production generates a shallow U-shaped curve for hours devoted to market work, and that it leads to the marketization of home production, i.e., the movement of time out of home production and into market production of services. Both of these patterns are found in the US data. The initial decrease in market work is associated with the movement of activity into services, which have better home produced substitutes. But as time advances, a higher rate of growth in the productivity of market produced services relative to home produced services leads to the movement of activity out of the home sector and into the market sector, which results in the increase in market hours.

Another dramatic trend in labor market outcomes has been the rise of female labor force participation. Several authors have argued that the process of structural transformation is an important factor in accounting for this change. The basic idea is that jobs in the goods sector (i.e., agriculture and manufacturing) and the service sector tend to have different weights on various dimensions of labor input. In particular, the goods sector places more emphasis on “brawn” while the service sector places more emphasis on “brains”. If men and women have different relative endowments of these two factors, then the movement of activity from one sector to the other could plausibly impact on the desire of women to seek employment in the market sector. Fuchs (1968) noted this explanation for the rise of female labor force participation. In recent work, Rendall (2010) builds a two-sector model in which she can quantitatively evaluate this role and argues that structural transformation is an important quantitative factor in accounting for the rise of female labor force participation. In related work, Akbulut (2011) also argues that

the rise of the service sector has been an important factor in accounting for the rise of female labor force participation in the US, but the key reallocation in her model is the movement of labor out of home produced services and into market produced services in response to a more rapid rate of technological progress in market services relative to home produced services.

## **7.5 Structural Transformation and Business Cycles**

There are many different ways in which theories of structural transformation and business cycles might overlap. One idea which frequently recurs is that some business cycles are the result of periods of greater reallocation of economic activity across sectors. To the extent that this reallocation of activity occurs at the broad sectoral level emphasized by models of structural transformation, structural transformation and business cycles could be intimately related.

Using the search model of Lucas and Prescott (1974) as a reference point, Lilien (1982) argued that if it takes time for labor to move from one sector to another, then periods of above average reallocation will also be periods of above average unemployment. He then argued that business cycles in the post World War II US were characterized as periods of above average reallocation of labor among two-digit sectors, as measured by the variance in employment growth rates at the two sector level. However, subsequent work by Abraham and Katz (1986) argued that Lilien's statistical finding about changes in the variance of sectoral growth rates could simply be due to the fact that sectors vary in their response to aggregate shocks, and that data on vacancies supported this latter explanation over the sectoral shifts explanation.

This idea has experienced a recent resurgence in popularity in the face of the current recession, with various economists suggesting that "mismatch" is an important element of the current high level of unemployment, and that the decline of broad sectors such as manufacturing and construction is an important element of this mismatch. However, despite its popularity, recent empirical research by Sahin et al. (2011) and Herz and van Rens (2011) finds little evidence for this explanation.

We note that even if reallocation were concentrated during recessions, it would not follow that recessions are caused by the reallocation. Rather, it may be that recessions are caused by a second factor, and that the decisions that lead to reallocation are made in such a way that

reallocation coincides with the recession. That is, for example, it may be that steel mills go out of business permanently during recessions, but this may simply reflect that the optimal timing of exit for a steel plant is during a downturn in economic activity. Rogerson (1991) argued that movement out of agriculture in the US has been concentrated during upturns in economic activity, whereas the movement of workers out of manufacturing has been concentrated during downturns.

Even if structural transformation is not the cause of business cycles, it may still exert an influence on business cycles. For example, to the extent that value added varies in volatility across sectors, the sectoral composition of aggregate output is a potentially important determinant of business cycle fluctuations. In what follows, we mention two examples of this idea.

The first example is Da Rocha and Restuccia (2006), who disaggregated the economy into agriculture and non-agriculture and documented that indeed there are important differences between the two sectors. In particular, they found that the agricultural sector is more volatile than the rest of the economy, is not correlated with the rest of the economy, and has counter-cyclical employment. They showed that this implies that countries with a larger agricultural sector have more volatile aggregate output and less volatile employment. Moreover, it implies that as structural transformations out of agriculture occur, business cycle properties across countries converge.

The second example of how the sectoral composition matters is due to Moro (2009) and Carvalho and Gabaix (2010). They disaggregated the economy into services and manufacturing, largely ignoring agriculture. They documented that the volatility of services is lower than in manufacturing. Moro (2009) argued that the reason for this is that the share of intermediate inputs is larger in manufacturing than in services. Irrespective of why the volatilities differ between the two sectors, the implication is that the volatility of aggregate output declines as the share of services increases along the path of structural transformation. Carvalho and Gabaix (2010) found that this accounts for most of the “great moderation” and its recent undoing. In particular, the great moderation is due to a decreasing share of manufacturing between 1975 and 1985 and its recent undoing in the form of rising aggregate volatility is due to the increase of the size of the financial sector.

## 7.6 Structural Transformation and Wage Inequality

One of the dramatic secular changes in the US economy over the last fifty years has been the marked increase in wage inequality that is associated with the return to skill. In a recent paper, Buera and Kaboski (2011) argue that this rising return to skill is intimately connected to the structural transformation of economic activity towards services. They document in time series data the same threshold behavior of value added in services that we have found above, that is, there is a threshold for per capita income at which point one observes an acceleration in the increase in the value added share for services. Interestingly, at that threshold there is also an increase in the fraction of the workforce that becomes skilled and the skill premium. In the context of the US they also document that the entire rise in the service sector's share of value added in the last fifty years is accounted for by growth in sub-sectors that have higher than average shares of skilled labor. They go on to build a model that links these patterns as the outcome of structural transformation that is driven by neutral productivity growth.

Income effects play a key role in their model. Their preference structure is similar to the one used by Foellmi and Zweimüller (2008) except that they assume each want is indivisible. We previously discussed how this leads to income effects. Wants can be satisfied either via home production or market production. Production in both sectors uses goods and labor. There are two types of labor: skilled and unskilled. Skilled labor is specialized to a particular want, is costly to acquire and is subject to an increasing cost curve. To capture the fact that home production is necessarily less specialized, they assume that skilled labor is equivalent to unskilled labor in home production. Wants differ in "complexity", where complexity captures both the amount of labor that is required to produce them and the relative productivity advantage of skilled labor in producing the want.

Unlike the work of Foellmi and Zweimüller (2008), these authors provide a definitive mapping between their model and sectoral data. Specifically, they assume that all of the wants in preferences represent service flows. If a flow is satisfied through the market then it counts toward the market service sector. If it is satisfied via home production, then it is not counted in market output. The goods sector in their model produces the goods that are used to produce service flows both in the home and the market. For example, production of motor vehicles counts

in the market production of goods, but if the motor vehicle is a private passenger car then the service flow of transportation services will not count as market production of services, but if the motor vehicle is a bus used by a public or private entity then the transportation services will count in market services.

As the economy develops it produces wants that are increasingly complex, thereby creating additional incentives for both market production of wants and skill accumulation. Because there is an upward sloping supply curve for skilled workers, the skill premium is also increasing. The structure of their model is such that the relative advantage of skilled labor in producing more complex goods only emerges beyond a critical threshold level of complexity, so that these patterns also emerge beyond a threshold. A key fact that this model is able to account for that our benchmark model cannot is that this model predicts that the share of services in nominal value added is flat below some threshold.

A key implication of this work is that adding human capital into the analysis and understanding the different role of human capital in various activities is an important ingredient in understanding some key features of structural transformation.

## **8 Conclusion**

Rather than attempt to summarize all of the material that we have covered, we would like to use this section to make a few observations about the previous literature and outline what we think the priorities for future research should be in this field. First, while the search for specifications that can simultaneously yield structural transformation and balanced growth have proven to be useful in organizing research, we believe that exact balanced growth should not be imposed as a requirement moving forward. The literature should instead focus on building models that can quantitatively account for the properties of structural transformation and in the process assess the importance of various economic mechanisms.

Second, as more economies become developed and more of their activity is in the service sector, it is probably of interest to consider a finer disaggregation that reflects the very different nature of the activities that currently get grouped together in services. For example, education

and health care are presumably very different activities than retail trade. Education and health both reflect investment activity, and they tend to use very different skill intensities for the labor that they employ. The work by Buera-Kaboski (2011) that we described is a first step in this direction. Third, the role of openness and how globalization is influencing the nature of structural transformation remains relatively unexplored. Fourth, there is a need for better data to assess the nature of structural transformation in economies that are currently relatively less developed and examine the extent to which their behavior differs from that of economies that are now rich. This may provide more information about what factors are responsible for the lack of development in some economies.

# Appendix A: Data Sources and Sector Assignments

## Historical data 1800–2008

- Data source: GDP per capita at international dollars
  - Data on GDP per capita at 1990 international dollars are from Maddison (2010) for all countries and most years. There are some years in the early 19th century for Belgium, Netherlands, Sweden, United Kingdom, and the United States when there are data on value added and employment shares, but Maddison does not report data on GDP per capita. We calculated GDP per capita at international dollars for these years in the following way. From alternative sources, we first calculated real GDP per capita for the missing years and for the first year for which Maddison’s data is available. We then calculated the growth rates between the missing years and the first year for which the Maddison data is available. Lastly, we combined the growth rates with the Maddison’s data to calculate the per capita GDP at international dollars for the missing years. Next we list the data sources for these calculations.
    1. *Belgium*. 1835–1845: real GDP from Groningen Growth and Development Centre, Historical National Accounts Database 2009, and population from Maddison (2010).
    2. *Netherlands*. 1807–1830: real GDP per capita from Smits et al. (2007).
    3. *Sweden*. 1800–1820: real GDP per capita from Krantz and Schön (2007).
    4. *United Kingdom*. 1800–1830: real GDP per capita from Clark (2009).
    5. *United States*. Louis Johnston and Samuel H. Williamson, “What Was the U.S. GDP Then?” MeasuringWorth, 2011.
- Data source: Value added at current prices
  - *Belgium*. 1835–1990: Groningen Growth and Development Centre, Historical National Accounts Database 2009. 1991–2007: EUKLEMS 2009.

- *Spain*. 1885–1940: Groningen Growth and Development Centre, Historical National Accounts Database 2009, 1953–2004: Groningen Growth and Development Centre 10–sector Database 2007.
  - *Finland*. 1860–2001: Groningen Growth and Development Centre, Historical National Accounts Database 2009.
  - *France*. 1815–1938: Groningen Growth and Development Centre, Historical National Accounts Database 2009, 1950–1960: Mitchell (2007) Table J2, 1970–2005: Groningen Growth and Development Centre 10–sector Database 2007.
  - *Japan*. 1885–1940: Groningen Growth and Development Centre, Historical National Accounts Database 2009, 1953–2004: Groningen Growth and Development Centre 10-sector Database 2007.
  - *Korea*. 1911–1940: Groningen Growth and Development Centre, Historical National Accounts Database 2009, 1953–2005: Groningen Growth and Development Centre 10–sector Database 2007.
  - *Netherlands*. 1807–1913: Smits et al. (2007), 1970–2005: Groningen Growth and Development Centre 10–sector database, August 2008.
  - *Sweden*. 1800–2000: Krantz and Schön (2007), 2000–2005: Groningen Growth and Development Centre 10-sector Database, August 2008.
  - *United Kingdom*. 1801, 1941–1851: Broadberry et al. (2011) Table 8–9, 1811–1831, 1860–1910, 1950: Mitchell (2007) Table J2, 1920–1938: Feinstein (1972) Table 9, 1960–2005: Groningen Growth and Development Centre 10–sector Database 2007.
  - *United States*. 1800–1900: Agriculture and Manufacturing, Gallman (1960), Services, Gallman and Weiss (1969), 1909–1918: King (1930), 1919–1928: Kuznets et al. (1941), 1929–1946: Carter et al., eds (2006) Table Ca35–53, 1947–2008: Value Added by Industry, Gross Domestic Product by Industry Accounts, Bureau of Economic Analysis.
- Data source: Employment



- *Belgium*. 1846–1961: Mitchell (2007) Table B1, 1970–2007: EUKLEMS 2009.
  - *Spain*. 1860–1964: Mitchell (2007) Table B1, 1970–2007: EUKLEMS 2009.
  - *Finland*. 1805–1960: Mitchell (2007) Table B1, 1970–2007: EUKLEMS 2009.
  - *France*. 1856–1968: Mitchell (2007) Table B1, 1970–2007: EUKLEMS 2009.
  - *Korea*. 1953–2005: Groningen Growth and Development Centre 10–sector Database 2007.
  - *Netherlands*. 1807–1913: Smits et al. (2007), 1920–1947: Mitchell (2007) Table B1, 1970–2005: Groningen Growth and Development Centre 10–sector Database 2008.
  - *Sweden*. 1850–2000: Krantz and Schön (2007), 2000–2005: Groningen Growth and Development Centre 10–sector Database 2008.
  - *United Kingdom*. 1801, 1813–1820 average assigned to 1817, 1851: Broadberry et al. (2011) Table 1 and Table 12, 1841: Mitchell (2007) Table B1, 1861–1938: Feinstein (1972) Table 59–60, 1948–2005: Groningen Growth and Development Centre 10–sector Database 2007.
  - *United States*. 1840–1920: Carter et al., eds (2006) Table Ba814–830, 1929–2008: NIPA Table 6.8 Persons Engaged in Production, Bureau of Economic Analysis.
- Sector assignments
    1. Agriculture corresponds to the sum of International Standard Industrial Classification (ISIC) sections A–B. If ISIC classification was not available, we assigned industries to agriculture if the source table heading said “Agriculture” or “Agriculture, forestry and fishing”
    2. Manufacturing corresponds to the sum of ISIC sections C, D, F and includes mining, manufacturing, and construction. If ISIC classification was not available, we assigned industries to manufacturing if the source table heading said “Mining” or “Extractive industries” or “Manufacturing” or “Construction” or “Electricity, Gas and Water Supply” or “Utilities”.

3. Services correspond to the sum of ISIC sections E, G–P and include utilities, wholesale, retail trade, hotels and restaurants, transport, storage and communication, finance, insurance, real estate, business services, and community social and personal services. If ISIC classification was not available, we assigned industries to services if the source table heading said “Commerce” or “Finance” or “Trade” or “Transport” or “Communication” or “Services”

## **EUKLEMS 2009**

- Data sources (EUKLEMS series code in brackets)
  1. Employment
    - Total hours worked by persons engaged in millions (H.EMP)
  2. Value added
    - Gross value added at current basic prices (VA)
- Sector assignment
  1. Agriculture corresponds to the sum of International Standard Industrial Classification (ISIC) sections A–B.
  2. Manufacturing corresponds to the sum of ISIC sections C, D, F and includes mining, manufacturing, and construction.
  3. Services correspond to the sum of ISIC sections E, G–P and includes utilities, wholesale, retail trade, hotels and restaurants, transport, storage and communication, finance, insurance, real estate, business services, and community social and personal services.

## **World Development Indicators 2010**

- Data sources (WDI series code in brackets)
  1. Employment

- Employment in agriculture (% of total employment) (SL.AGR.EMPL.ZS)
  - Employment in industry (% of total employment) (SL.IND.EMPL.ZS)
  - Employment in services (% of total employment) (SL.SRV.EMPL.ZS)
2. Value added
- Agriculture, value added as % of GDP (NV.AGR.TOTL.ZS)
  - Industry, value added as % of GDP (NV.IND.TOTL.ZS)
  - Services, etc., value added as % of GDP (NV.SRV.TETC.ZS)
- Oil production
    1. Oil rents as % of GDP, (NY.GDP.PETR.RT.ZS)
  - Sector assignment
    1. Agriculture corresponds to the sum of ISIC divisions 1–5 and includes forestry, hunting, and fishing, as well as the cultivation of crops and livestock production.
    2. Manufacturing corresponds to the category “Industry” in the WDI, which is the sum of ISIC divisions 10–45 and includes mining, manufacturing, construction, electricity, water, and gas.
    3. Services correspond to the sum of ISIC divisions 50–99 and include value added in wholesale and retail trade (including hotels and restaurants), transport, and government, financial, professional, and personal services (such as education), health care, and real estate services. They also include imputed bank service charges, import duties, and statistical discrepancies as well as discrepancies arising from rescaling.

## **National Accounts of the United Nations Statistics Division**

- Data sources
  1. Gross value added by economic activity at current prices in national currency
- Sector assignment

1. Agriculture corresponds ISIC sections A–B.
2. Manufacturing corresponds to the sum of ISIC sections C–F and includes mining, manufacturing, utilities, and construction.
3. Services correspond to the sum of ISIC sections G–P and includes wholesale, retail trade, hotels and restaurants, transport, storage and communication, finance, insurance, real estate, business services, and community social and personal services.

## **Historical Consumption Shares UK and US**

- Data source: GDP per capita at international dollars at 1990 international dollars are from Maddison (2010)
- Data source: US Consumption share in current prices
  - 1900–1928: Carter et al., eds (2006)
  - 1929–2008: BEA
- Data source: UK Consumption share in current prices
  - 1900–1964: Feinstein (1972)
  - 1965–2008: Office of National Statistics (ONS)

## **Penn World Tables**

- Data source: PWT6.3 (PWT series code in brackets)
  1. Real Gross Domestic Product per Capita Relative to the United States (G-K method, current price) (y)
  2. Real GDP per capita in constant prices: chain series (rgdpch)
  3. Real GDP per worker in constant prices: chain series (rgdpwok)
  4. Population (pop)

- Data source: PWT benchmark 1980
  - Sector assignment
    1. Agriculture corresponds to the sum of PWT80 items 1–50
    2. Manufacturing corresponds to the sum of PWT80 items 51–54, 56–58, 63–66, 68–78, 81–83, 91–93, 95–97, 103–108, 112–113, 118–122
    3. Services correspond to the sum of PWT items 55, 59–62, 67, 79–80, 84–90, 94, 98–102, 109–111, 114–118, 123–125
  
- Data source: PWT benchmark 1985
  - Sector assignment
    1. Agriculture corresponds to the sum of PWT80 items 1–41
    2. Manufacturing corresponds to the sum of PWT80 items 42–47, 49–51, 56–61, 63–68, 70–72, 75–77, 82–84, 86–87, 94–97, 101, 107–109
    3. Services correspond to the sum of PWT items 48, 52–55, 62, 69, 73–74, 78–81, 85, 88–93, 98–100, 102–106
  
- Data source: PWT benchmark 1996
  - Sector assignment
    1. Agriculture corresponds to bread and cereals, meat, fish, milk, cheese and eggs, oils and fats, fruit, vegetables and potatoes, other food, non-alcoholic beverages, alcoholic beverages.
    2. Manufacturing corresponds to tobacco, clothing including repairs, footwear including repairs, fuel and power, furniture, floor coverings and repairs, other household goods incl. household textiles, household appliances and repairs, personal transportation equipment.
    3. Services correspond to gross rent and water charges, medical and health services, operation of transportation equipment, purchased transport services, com-

munication, recreation and culture, education, restaurants, cafes and hotels, other goods and services.

## **OECD Consumption Expenditure Data**

- Data source:
  - Final consumption expenditure of households, national currency, current prices, OECD National Accounts Statistics. This data set includes the final consumption expenditure of households broken down by the COICOP (Classification of Individual Consumption According to Purpose) classification and by durability.
- Sector assignment (COICOP codes in brackets)
  1. Food: “Food and non–alcoholic beverages” (P31CP010)
  2. Manufactured goods: “Durable goods” plus “Semi–durable goods” plus “Non–durable goods” minus “Food and non–alcoholic beverages” (P311B+P312B+P313B–P31CP010)
  3. Services: Services (P314B)
- Construction of the data for E7 countries (Austria, Denmark, Finland, France, Italy, Netherlands, United Kingdom) for the period 1980–2009. Consumption expenditure data are from the National Accounts of Eurostat both in local currency and euro. Then, for each year and each country, a conversion rate between local currency and euro was calculated by dividing total consumption expenditures in local currency with total consumption expenditures in euros. The three expenditure items expressed in local currency were converted into euros using this conversion rate, and then they were aggregated.

## **Real GDP per capita at 1990 international \$**

- Prior to 1970 the data on GDP per capita at 1990 international dollars are from Maddison (2010) for all years and countries if it was available,

- After 1970 we constructed real GDP per capita at 1990 international \$ in the following ways. The data on GDP per capita at 1990 international dollars for the United States were taken from Maddison (2010). The real GDP per capita of the United States was multiplied by the data on real GDP per capita relative to the United States to calculate the real GDP per capita at 1990 international \$ for each country and each year.

## Appendix B: Panel Regressions

To get a balanced panel, we only include countries with data over the entire period 1970–2007. In addition, we restrict the sample in three ways: we exclude countries in which the average ratio of oil rent to GDP exceeds 20% during 1970–2007;<sup>40</sup> we exclude countries with average populations of fewer than a million during 1970–2007; we exclude the former communist countries. The reason for these exclusion criteria is that the sector composition in these countries may be distorted. This leaves 103 countries.

**Table 1: Panel Data Analysis Agriculture, 1970–2007**

	Dependent variable: Agricultural share in value added					
	(1)	(2)	(3)	(4)	(5)	(6)
log GDP per capita	-0.121** (0.001)	-0.489** (0.021)	0.450* (0.184)	-0.126** (0.015)	-0.396** (0.067)	0.169 (0.274)
(log GDP per capita) <sup>2</sup>		0.022** (0.001)	-0.096** (0.022)		0.017** (0.004)	-0.056 (0.035)
(log GDP per capita) <sup>3</sup>			0.005** (0.001)			0.003* (0.001)
Country fixed effects	No	No	No	Yes	Yes	Yes
$R^2$	0.751	0.783	0.786	0.751	0.781	0.784
$N$	3914	3914	3914	3914	3914	3914

*Notes:* Heteroscedasticity robust standard errors in parentheses. Significance levels are indicated by †  $p < 0.10$ , \*  $p < 0.05$ , \*\*  $p < 0.01$ .

<sup>40</sup>The oil-rent-to-GDP ratio is taken from the WDI.



**Table 2: Panel Data Analysis Manufacturing, 1970–2007**

	Dependent variable: Manufacturing share in value added					
	(1)	(2)	(3)	(4)	(5)	(6)
log GDP per capita	0.043** (0.001)	0.447** (0.021)	-1.196** (0.144)	0.054** (0.017)	0.497** (0.078)	-1.252** (0.446)
(log GDP per capita) <sup>2</sup>		-0.025** (0.001)	0.182** (0.018)		-0.028** (0.005)	0.198** (0.058)
(log GDP per capita) <sup>3</sup>			-0.009** (0.001)			-0.009** (0.002)
$R^2$	0.234	0.331	0.352	0.234	0.331	0.348
$N$	3914	3914	3914	3914	3914	3914

*Notes:* Heteroscedasticity robust standard errors in parentheses. Significance levels are indicated by <sup>†</sup>  $p < 0.10$ , \*  $p < 0.05$ , \*\*  $p < 0.01$ .

**Table 3: Panel Data Analysis Services, 1970–2007**

	Dependent variable: Service share in value added					
	(1)	(2)	(3)	(4)	(5)	(6)
log GDP per capita	0.078** (0.001)	0.041* (0.019)	0.745** (0.170)	0.072** (0.012)	-0.101 (0.089)	1.084* (0.417)
(log GDP per capita) <sup>2</sup>		0.002* (0.001)	-0.086** (0.021)		0.011 <sup>†</sup> (0.006)	-0.142* (0.055)
(log GDP per capita) <sup>3</sup>			0.004** (0.001)			0.006** (0.002)
$R^2$	0.493	0.493	0.496	0.493	0.485	0.476
$N$	3914	3914	3914	3914	3914	3914

*Notes:* Heteroscedasticity robust standard errors in parentheses. Significance levels are indicated by <sup>†</sup>  $p < 0.10$ , \*  $p < 0.05$ , \*\*  $p < 0.01$ .

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